

The Political Legacy of American Slavery^{*}

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Abstract

We show that contemporary differences in political attitudes across counties in the American South in part trace their origins to slavery's prevalence more than 150 years ago. Whites who currently live in Southern counties that had high shares of slaves in 1860 are more likely to identify as a Republican, oppose affirmative action, and express racial resentment and colder feelings toward blacks. These results cannot be explained by existing theories, including the theory of contemporary racial threat. To explain these results, we offer evidence for a new theory involving the *historical persistence* of political and racial attitudes. Following the Civil War, Southern whites faced political and economic incentives to reinforce existing racist norms and institutions to maintain control over the newly free African-American population. This amplified local differences in racially conservative political attitudes, which in turn have been passed down locally across generations. Our results challenge the interpretation of a vast literature on racial attitudes in the American South.

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1 Introduction

For the first 250 years of American history, white landowners, predominantly from the South, enslaved millions of individuals of African descent. This “peculiar institution,” as it was sometimes called, defined the social, economic, and political landscape of the American South throughout this period. Slavery was so crucial to the South that one Georgia newspaper editor wrote, “negro slavery is the South, and the South is negro slavery” (cited in Faust, 1988). Yet, despite slavery’s prominence in shaping American history, and despite volumes written by economists and historians on its consequences, political scientists have largely overlooked how American slavery and the events following its abolition could continue to influence the South’s *contemporary* politics. Given recent findings on the long-term consequences of past events and institutions (Dell, 2010; Nunn and Wantchekon, 2011; Acemoglu, García-Jimeno and Robinson, 2012; Voigtländer and Voth, 2012; Alesina, Giuliano and Nunn, 2013), it would be surprising if such a fundamental aspect of American history had no persistent impact on American politics.

In this paper, we show that the local prevalence of slavery—an institution that was abolished 150 years ago—has a detectable effect on present-day political attitudes in the American South. Drawing on a sample of more than 40,000 Southern whites and historical census records, we show that whites who currently live in counties that had high concentrations of slaves in 1860 are today on average more conservative and express colder feelings toward African Americans than whites who live elsewhere in the South. That is, the larger the number of slaves per capita in his or her county of residence in 1860, the greater the probability that a white Southerner today will identify as a Republican, oppose affirmative action, and express attitudes indicating some level of “racial resentment.” We show that these differences are robust to accounting for a variety of factors, including geography and mid-19th century economic and social conditions. These results strengthen when we instrument for the prevalence of slavery using geographic variation in cotton growing conditions.

We consider several explanations for our results rooted in contemporary forces and find each to be inconsistent with the empirical evidence. For example, we consider the possibility that whites are simply more racially conservative when exposed to larger black populations—the central finding of the literature on racial threat (Key, 1949; Blalock, 1967; Blumer, 1958). However, when we estimate the direct effect of slavery on contemporary attitudes (Acharya, Blackwell and Sen, 2016), we find that

contemporary shares of the black population explain little of slavery's effects. We also test various other explanations, including the possibility that slavery's effects are driven exclusively by 20th-century population shifts or income inequality between African Americans and whites. We find no evidence that these contemporary factors and theories of population sorting fully account for our results. Introducing individual-level and contextual covariates commonly used in the public opinion literature also does not explain away our finding.

To explain our results, we instead propose a theory of the historical persistence of political attitudes. The evidence suggests that regional differences in contemporary white attitudes in part trace their origins to the late slave period and the time period after its collapse, with prior work suggesting that the fall of slavery was a cataclysmic event that undermined Southern whites' political and economic power. For example, [Key \(1949\)](#), [Du Bois \(1935\)](#), and [Foner \(2011\)](#) (among others) have argued that the sudden enfranchisement of blacks was politically threatening to whites, who for centuries had enjoyed exclusive political power. In addition, the emancipation of Southern slaves undermined whites' economic power by abruptly increasing black wages, raising labor costs, and threatening the viability of the Southern plantation economy ([Ransom and Sutch, 2001a](#); [Alston and Ferrie, 1993](#)). Taken in tandem with massive preexisting racial hostility throughout the South, these political and economic changes gave Southern Black Belt elites an incentive to further promote existing anti-black sentiment in their local communities by encouraging violence towards blacks and racist attitudes and policies ([Roithmayr, 2010](#)). This amplified the differences in white racial hostility between former slaveholding areas and nonslaveholding areas, and intensified racially conservative political attitudes within the Black Belt. These have been passed down locally, one generation to the next.

We provide empirical support for this mechanism by showing that areas of the South that were the earliest to eliminate the political and economic incentives for anti-black violence—for example, by adopting new technologies, such as tractors, that reduced the demand for black farm labor—are also the areas in which slavery's long term effects have most attenuated. Furthermore, as evidence for intergenerational (cultural) transfer of attitudes, we show that there exists a strong correlation between the racial attitudes of parents and their children in the U.S. South. Our evidence, therefore, supports the theory that political attitudes have persisted historically in the U.S. South, rather than the view that attitudes are driven exclusively by contemporaneous forces—making our position quite distinct from much of the existing public opinion literature.

The paper proceeds as follows. In Section 2, we motivate our hypothesis that historical persistence—and not just contemporary factors—shape modern-day political attitudes. We discuss our data in Section 3 and present our core results linking the prevalence of slavery in 1860 and contemporary attitudes in Section 4, with additional robustness checks presented in the Supplemental Information. In Section 5, we consider and provide evidence against several competing theories rooted in contemporary factors, including the theory of racial threat. In Section 6, we provide evidence for our theory of the historical persistence of political attitudes, paying close attention to postbellum political and economic incentives as the driving mechanism. Section 7 concludes by discussing the broader implications of our research for scholarship in American political behavior.

2 Explaining Regional Differences in Southern Political and Racial Attitudes

We orient our analysis toward the Southern “Black Belt” (or the “Cotton Belt”), the hook-shaped swath of land that was the primary locus of antebellum slavery (Figure 1). Scholars have noted that Black Belt whites were particularly prominent in Southern politics and have been more conservative than whites elsewhere in the South. As V.O. Key wrote, it is “the whites of the black belts who have the deepest and most immediate concern about the maintenance of white supremacy,” and “if the politics of the South revolves around any single theme, it is that of the role of the black belts” (Key, 1949). Furthermore, the Black Belt has had an enormous influence on national politics. Members of Congress from these areas held influential positions, effectively exercising veto power during the development of the welfare state in the 1920s and 30s (Katznelson, Geiger and Kryder, 1993). Given these facts, our motivating question is: Why are whites who currently live in the Black Belt more conservative than whites living elsewhere in the South, particularly on race-related issues? We consider two broad classes of explanations: (1) the historical persistence of attitudes originating in slavery and (2) contemporary factors, including contemporary demographics and geographic mobility.

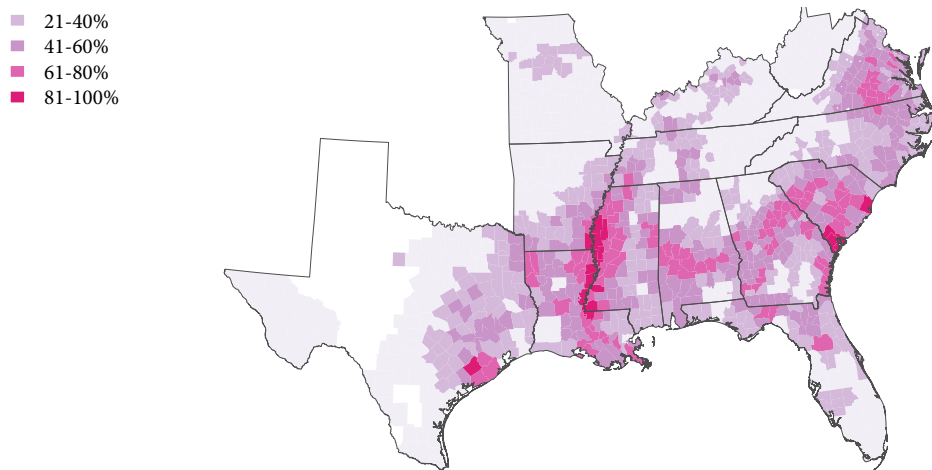


Figure 1: Estimated proportion slave in 1860 by county.

2.1 Historical Persistence of White Political Attitudes

Our first hypothesis is that today’s Black Belt is more politically conservative than other parts of the South in part because of its history of chattel slavery. We are motivated in this hypothesis by an emerging empirical literature showing that the effects of coercive institutions persist in other contexts. [Dell \(2010\)](#), for example, shows that a colonial forced labor system in Peru and Bolivia led to lower levels of modern-day household consumption and childhood growth. [Acemoglu, García-Jimeno and Robinson \(2012\)](#) find that the use of slaves in the colonial gold mines of Colombia predicts modern-day poverty, reduced school enrollment, and decreased vaccination rates. [Nunn and Wantchekon \(2011\)](#) show that Africans whose ancestors were targeted by the slave trade have higher levels of mistrust today than other Africans. Within the United States, [O’Connell \(2012\)](#) demonstrates that areas of the American South that had high numbers of slaves have greater economic inequality between blacks and whites today. Similarly, [Nunn \(2008\)](#) and [Lagerlöf \(2005\)](#) find a negative relationship between the prevalence of slavery and income in the American South, and [Mitchener and McLean \(2003\)](#) find a negative relationship between slavery and modern-day labor productivity. These papers are part of a growing literature that shows that historical institutions such as slavery can affect both institutional and behavioral outcomes long after the institutions themselves disappear ([Nunn, 2009](#)). This work complements an existing literature documenting the path dependence of historical institutions over time (e.g., [Pierson, 2000, 2004](#)).

Building from this literature, we hypothesize that Southern slavery may have had a similarly lasting effect on political and racial attitudes. The rise and swift fall of chattel slavery together were cataclysmic events. Specifically, the eventual fall of slavery undermined the political and economic power of the Southern whites, particularly in the Black Belt, (Du Bois, 1935; Foner, 2011), making them more hostile toward African Americans and conservative in their political, racial, and economic views (noted contemporaneously by Key, 1949). Qualitative accounts (as we document in Section 6) suggest that the nature of Southern whites' responses to the collapse of slavery varied according to how locally prevalent—and thus politically and economically important—slavery had been. Areas with more enslaved people reacted more sharply to emancipation by curtailing blacks' rights and oppressing newly freedmen and their mobility.

In addition, a large literature has shown that attitudes can persist historically through both cultural and institutional channels (Nunn and Wantchekon, 2011; Voigtländer and Voth, 2012). On the one hand, Southern institutions such as Jim Crow helped enforce racial segregation, while racially targeted violence reinforced practices of black subjugation (Woodward, 2002 [1955]). On the other hand, the culture of the Southern Black Belt was one where black subjugation was passed on within white families and across generations—a process that no doubt included intergenerational socialization (Boyd and Richerson, 1988; Bisin and Verdier, 2000; Campbell et al., 1980; Jennings and Niemi, 1968).

Based on these arguments, we expect that: (1) areas that were more reliant on slavery should be more conservative today on race-related issues and in terms of party affiliation, (2) race-related attitudes should be correlated across generations within the South, and (3) the effects of slavery should be weaker (that is, should have decayed more) in areas where the incentives for anti-black attitudes faded earlier.

2.2 How Contemporary Demographics Could Explain Regional Variation in White Political Attitudes

In contrast to the arguments above, much of the political science literature points to contemporary (not historical) forces as providing the explanation for why Black Belt whites are more conservative on race. By and large, the literature has interpreted Key (1949)'s work as suggesting that whites contemporaneously become more conservative when they are exposed to the high concentrations of African Americans who live

in their communities.¹ The high concentration of African Americans in today’s Black Belt could contemporaneously threaten white dominance, resulting in whites actively choosing more conservative political beliefs today. The literature supporting this idea, known as “racial threat,” is voluminous.² For example, [Glaser \(1994\)](#) finds evidence linking negative white attitudes toward civil rights or African-American politicians with high concentrations of blacks. [Giles and Buckner \(1993\)](#) find a relationship between black concentrations and white support for racially conservative candidates such as David Duke (these findings are, however, challenged by [Voss, 1996](#)). This literature, however, has not considered that slavery could be an *independent* predictor of contemporary attitudes (apart from its effect on contemporary demographics), making it an omitted variable in studies of racial threat in the South.

Other aspects of the contemporary local context may also affect white attitudes—for example, income gaps between blacks and whites, urban-rural differences, and other contextual and individual-level factors (e.g., [Oliver and Mendelberg, 2000](#); [Hopkins, 2010](#); [Verba, Schlozman and Brady, 1995](#)).³ A final category of explanations concern white mobility through the 20th century. For example, it could be that more racially conservative whites have migrated into former slaveholding areas, while racial liberals have left, thereby creating a regional pattern in political attitudes that is less about persistence of beliefs and more about the sorting of beliefs.

Ultimately, much of the public opinion literature focuses on contemporary or individual-level factors in explaining political beliefs, rather than on historical forces. Yet, [Key](#) himself was aware of the importance of history in the context of slavery when he noted that, in the years leading to the Civil War, “those with most at stake—the owners of large numbers of slaves—were to be found roughly in the same areas as present-day black belts” ([Key, 1949](#)). We now turn to exploring this historical link in terms of regional variation in Southern whites’ attitudes.

¹As we note in Section 5.1, Southern slavery is correlated with contemporary black concentration, making it difficult to disentangle the effects of slavery from the effects of contemporary black concentrations.

²Early studies showed, for example, that modern black concentrations predict white support for segregationist candidates such as George Wallace (e.g., [Wright Jr, 1977](#); [Black and Black, 1973](#)), racially hostile white attitudes ([Giles, 1977](#); [Blalock, 1967](#)), negative attitudes on school desegregation ([Ogburn and Grigg, 1956](#)), and higher incidence of lynchings ([Reed, 1972](#)).

³Some work has even highlighted the connection between slavery and these contemporary factors ([O’Connell, 2012](#); [Nunn, 2008](#); [Mitchener and McLean, 2003](#)). While these papers suggest that slavery might affect contemporary attitudes indirectly through contemporary factors such as economic inequality and prosperity, we find in Section 5 that slavery has a *direct* effect on contemporary attitudes that does not work through these channels.

3 Historical Slave Data and Contemporary Public Opinion Data

Our main explanatory variable and proxy for slavery's prevalence is the proportion of each county's 1860 population that was enslaved, as measured by the 1860 U.S. Census. Although counts of enslaved people were taken before 1860, we use measures from 1860 because they represent the last record before chattel slavery was abolished in 1865. In addition, white planters were very mobile in the antebellum period, during which slaves (not land) were their main source of wealth; after emancipation, mobility decreased rapidly as white elites became increasingly oriented toward landowning (Wright, 1986, 34). If any local legacy exists, we would expect to see it in data from 1860. Since county boundaries have shifted since 1860, we use an area-weighting method to map data from the 1860 Census onto county boundaries in 2000, enabling us to estimate the proportion enslaved in 1860 within modern-day counties.⁴ Figure 1 depicts the data. Overall, we have in our data approximately four million enslaved people, constituting 32% of the Southern population.

3.1 Outcome Variables Measuring Contemporary White Political and Racial Attitudes

We analyze three county-level outcome measures, which come from the Cooperative Congressional Election Study (CCES), a large survey of American adults (Ansolabehere, 2010). We pool CCES data from the 2006, 2008, 2009, 2010, and 2011 surveys to create a combined data set of over 157,000 respondents. We subset these data to the former Confederate States plus Missouri and Kentucky, both of which had significant internal support for the Confederacy,⁵ and to self-identified whites, leaving us with more than 40,000 respondents across 1,329 of the 1,435 Southern counties. In addition, we also investigate individual-level black-white thermometer scores from waves of the American National Election Survey (ANES) from 1984 until 1998, a time period where the ANES both used a consistent sampling frame and included county-level identifiers for

⁴Total population and total enslaved population in 1860 counties are divided among the counties in 2000 so that the proportion of the 1860 population from 1860-county i that is assigned to 2000-county j is based on the size of their overlapping areas. This approach produces estimates and results similar to those provided by (1) O'Connell (2012) ($r = 0.986$), who uses an alternative interpolation technique, and (2) a simpler method that relies on matching counties by name. See Appendix A for more information on our approach.

⁵The sample thus includes AL, AR, FL, GA, KY, LA, MS, MO, NC, SC, TN, TX, VA, and WV.

respondents. After restricting the sample to Southern whites, we have an ANES sample of 3,123 individuals across 64 counties in the South. This makes the ANES more restricted in its geographic coverage, but it contains valuable direct questions on the subjective evaluation of racial groups.

The four outcome measures are as follows.

Partisanship. We examine partisanship because, as many scholars have argued, Southern whites' partisanship (and partisan re-alignment) has been intimately connected to, and reflective of, their attitudes on race and black-white relations (Key, 1949; Carmines and Stimson, 1989; Valentino and Sears, 2005; Kuziemko and Washington, N.d.). Such partisan identification can not only reflect racial attitudes, as suggested by these papers, but may also reflect beliefs on policy issues closely related to race, including redistribution (Gilens, 2009; Lee and Roemer, 2006). Partisanship also serves as an important bridge to regional and national politics.

We construct our partisanship measure from a standard seven-point party identification question on the CCES. We operationalize the party variable as whether an individual identified at all with the Democratic Party (1 if Democrat; 0 otherwise).⁶ Thus the county-level measure represents the proportion of whites in each county who identified as Democrats.

Support for affirmative action. All CCES surveys ask respondents whether they support or oppose affirmative action policies, which are described as "programs [that] give preference to racial minorities and to women in employment and college admissions in order to correct for discrimination" (2008 CCES).⁷ We construct the outcome variable by collapsing the four-point scale, from "strongly support" to "strongly oppose," to an indicator representing whether the respondent demonstrated any level of support for affirmative action (1 for support; 0 otherwise). At the county level, then, this is the proportion of whites who say that they support affirmative action.

Racial resentment. Kinder and Sears (1981) write that racial resentment (or symbolic racism) "represents a form of resistance to change in the racial status quo based on

⁶We use survey data as opposed to voter registration data because primaries in many Southern states are open. Coupled with the dramatic changes in partisanship in the South over the last 40 years, this means voter registration data are unreliable measures of current partisan leanings. Finally, survey data allows us to focus on the partisanship of whites voters only.

⁷Although the question wording differs across years, we have no reason to believe that these wording variations affect our analysis.

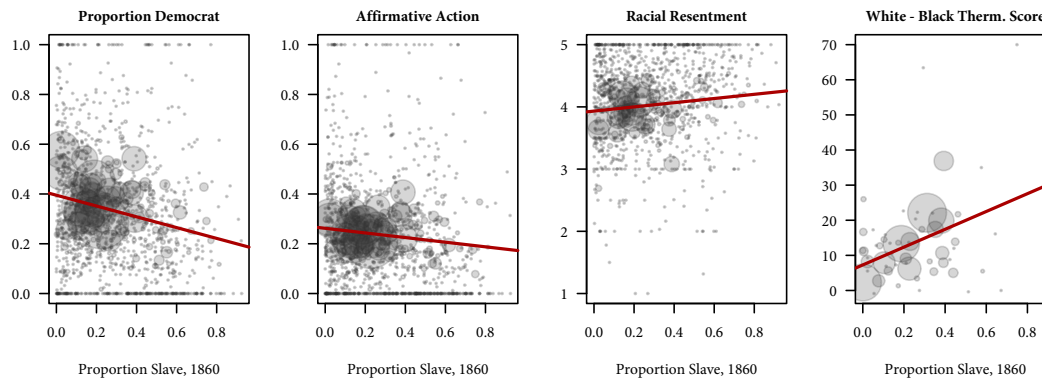


Figure 2: Bivariate relationships between proportion slave in 1860 and the four outcome measures with a linear fit in red. All four relationships are significant at $p < 0.05$ significance levels. Size of the points are in proportion to their within-county sample size (weighted by sampling weights).

moral feelings that blacks violate such traditional American values as individualism and self-reliance, the work ethic, obedience, and discipline.” We construct a third outcome variable using the two CCES questions on racial resentment. The first question, asked in the 2010 and 2011 CCES surveys, asks respondents on a five-point scale whether they agree with the following statement: “The Irish, Italian, Jews and many other minorities overcame prejudice and worked their way up. Blacks should do the same.” The second question, asked in 2010, asks respondents, also on a five-point scale, whether they agree that “Generations of slavery and discrimination have created conditions that make it difficult for Blacks to work their way out of the lower class.” For the 2010 CCES, when both questions were asked, we rescaled both questions and averaged them to create one measure. The final county-level measure is the average level of agreement with the racially resentment statement on a five-point scale.

White-black thermometer difference. In many years, the ANES contains “feeling thermometer” questions, which ask respondents to evaluate their feelings about politicians and groups (including racial or ethnic groups) on a scale from 0 to 100.⁸ Since

⁸The 1984 ANES gave respondents the following instructions:

I’ll read the name of a person and I’d like you to rate the person using the feeling thermometer. Ratings between 50 degrees and 100 degrees mean that you feel favorable and warm toward the person. Ratings between 0 degrees and 50 degrees mean that you don’t feel favorable toward the person and that you don’t care too much for that person. You would rate the person at the 50 degree mark if you don’t feel particularly warm or cold toward that person.

these scales have engendered criticisms that the ratings fluctuate heavily from individual to individual (Wilcox, Sigelman and Cook, 1989) and that they are less stable than party identification (Markus and Converse, 1979), we use them mainly to buttress the findings from our previous three outcome variables. To measure relative racial hostility, we take the difference between white respondents' feeling thermometer ratings toward whites and their feeling thermometer ratings towards blacks. Thus, a positive difference would indicate that respondents have warmer feelings towards whites as opposed to blacks. Only using black thermometer scores yields similar results, but we use the difference in case slavery has an overall effect on racial group thermometer ratings.

Appendix Tables A.1 and A.2 report summary statistics for these and other data. Figure 2 depicts the bivariate relationships between proportion slave in 1860 and the four outcome measures from the CCES and ANES. It shows negative, statistically significant relationships between slave prevalence and proportion Democrat and support for affirmative action and positive, statistically significant relationships with racial resentment and thermometer score differences. We now turn to establishing plausible estimates of the causal effect of slavery's prevalence on these outcomes.

4 Slavery's Effects on Contemporary Outcomes

In models that analyze the CCES outcomes, we conduct a county-level analysis using weighted least squares (WLS) with the within-county sample size (appropriately weighted by the sampling weights) as weights.⁹ We opt for analyses at the county level for two reasons. First, the key variable of interest here (in effect, our treatment) is slavery in 1860, which is only measured at the county level, meaning we cannot detect any within-county variation in this measure. Second, when attempting to estimate the effect of a cluster-level variable such as this, Green and Vavreck (2008) show that aggregating units up to the cluster level provides the most accurate variance estimates. While it is sometimes useful to trade off this accuracy on standard error estimation to allow for the inclusion of individual-level covariates, in our context many

For groups like "blacks" or "whites," the instructions asked "And, still using the thermometer, how would you rate the following?"

⁹Whites in high-black areas are less likely to be sampled than whites in low-black areas, since the sampling reflects the population distribution of whites. If we were making inferences about both blacks and whites, this approach would be problematic. However, we use the white population subset to make inferences only about the Southern white population.

individual-level covariates are potentially directly affected by the treatment (slavery) so including them may introduce post-treatment bias (Rosenbaum, 1984). In spite of this, we present respondent-level analyses in Appendix Table A.3 that include additional respondent-level controls and standard errors clustered at the county level. These results are consistent with our county-level results below, although we caution that the estimates may be biased. Furthermore, we opt for an individual-level analysis for our results on NES thermometer scores since we are collecting together 15 years' worth of survey results wherein racial attitudes have changed dramatically, making it important to include survey-year fixed effects in our models. The level of analysis has very little impact on the substantive results, and the ANES results at the county level are presented in Appendix Table A.11.

In Table 1, we report the baseline estimates of slavery's effect on the three CCES outcomes. Column (1) of Table 1 presents the simple WLS relationship between slavery and white partisan identification, which measures the relationship depicted in Figure 2. In the remaining columns, we include state-level fixed effects to address the possibility that states adopted different policies that could have influenced slave shares in 1860 and could affect our outcome variables in ways unrelated to slavery. In addition, in the remaining columns we control for factors that may have been predictive of proportion slave in 1860. These "1860 covariates," unless otherwise noted, come from the 1860 U.S. Census and address possible differences between slaveholding and non-slaveholding counties. First, since wealthier or more populous counties may have had more or fewer slaves, we control for economic and demographic indicators from 1860. These include (i) the log of the total county population, (ii) the proportion of farms in the county smaller than 50 acres, (iii) the inequality of farmland holdings as measured by the Gini coefficient for landownership (Nunn, 2008), (iv) the log of total farm value per improved acre of farmland in the county, and (v) the log of the acres of improved farmland. Second, because counties may have had different norms about race, we include controls for (vi) the proportion of total population in 1860 that is free black. We also control for characteristics related to trade and commerce, including separate indicators for whether the county had access to (vii) rails and (viii) steamboat-navigable rivers or canals. Finally, to account for any remaining spatial variation, we control for (ix) the log of the county acreage, (x) the ruggedness of the county terrain (Hornbeck and Naidu, 2014), and (xi) the latitude and longitude of the county, as well as their squared terms (to flexibly control for spatial variation in the outcome). For all of these variables, we mapped 1860 data onto modern county boundaries using the procedure

Table 1: Effects of slavery on white political attitudes.

	Proportion Democrat		Support for Affirm. Action	Racial Resentment	White-Black Therm. Diff
	(1)	(2)	(3)	(4)	(5)
Prop. Slave, 1860	-0.217** (0.023)	-0.127** (0.043)	-0.145** (0.036)	0.463** (0.139)	36.125** (9.211)
Level	County	County	County	County	Individual
1860 Covariates		✓	✓	✓	
State Fixed Effects		✓	✓	✓	✓
State-Year Fixed Effects					✓
Clustered SEs					✓
N	1,242	1,152	1,152	1,027	1,489
R ²	0.065	0.203	0.101	0.143	0.183

Notes: *p < .05; **p < .01. Models 1-4 are WLS at the county-level, with within-county sample sizes, adjusted by sampling weights, as weights. Model 5 is WLS at the individual-level with state-year fixed effects. Standard errors in parentheses, which are clustered at the county level in model 5. There are 50 counties in model 5.

described in Appendix A.¹⁰

As Columns (2) - (4) show, the conditional effects of slavery are meaningful and significant for all three CCES outcome variables. To illustrate, a 20 percentage-point increase in the slave proportion (roughly a one standard-deviation change) is associated with a 2.3 percentage-point decrease in the share of whites who currently identify as Democrats (and so roughly a 4.6 percentage-point shift toward the Republican Party), a 3 percentage-point decrease among those who currently support affirmative action, a 0.08 point increase in the average racial resentment score in the county, and a 7 point increase in the county average thermometer preference for whites relative to blacks. These represent approximately a 0.09-0.13 standard deviation change in the CCES outcomes and roughly a 0.32 standard deviation change in the thermometer score difference. The individual-level results for the CCES in Appendix Table A.3 give very similar estimates, making it unlikely that these results are due to some sort of ecological or aggregation bias.¹¹ Finally, we present results for each racial thermometer score sepa-

¹⁰In other results, we have additionally controlled for antebellum county-level vote share for the Democrats in various presidential elections. These controls don't affect the results but have the unfortunate byproduct of dropping South Carolina, which decided its electoral votes by the state legislature. For that reason, we do not include this particular control here.

¹¹In Appendix Table A.4, we present results with a full set of interactions between slavery and individual-level covariates and we find no significant interactions with education, household income, religiosity, gender, or age. In the same table, we show that there are no consistent interactions with contemporary contextual variables as well.

Table 2: Instrumental variables estimates of the effect of slavery.

	Proportion Slave, 1860 (1)	Proportion Democrat (2)	Support for Affirm. Action (3)	Racial Resentment (4)	White-Black Therm. Diff (5)
Cotton Suitability	0.441** (0.037)				
Prop. Slave, 1860		-0.277** (0.104)	-0.247** (0.088)	0.823* (0.352)	51.748* (19.305)
State Fixed Effects	✓	✓	✓	✓	✓
Geographic Controls	✓	✓	✓	✓	✓
Model	2SLS First Stage	2SLS Second Stage	2SLS Second Stage	2SLS Second Stage	2SLS Second Stage
N	1,120	1,120	1,120	998	51
F Statistic	80.077** (df = 21; 1098)				

Notes: † p < .1; * p < .05; ** p < .01. Model 1 is the first-stage relationship. Model 2-5 are weighted two-stage least-squares models, with weights as given in Table 1.

rately in Appendix Table A.10 and the full set of county-level results for thermometer scores in Appendix Table A.11, both of which are consistent with results in Table 1.

4.1 Instrumenting for Slavery with Cotton Suitability

There are two potential concerns with the above as a causal analysis. First, the 1860 slave data are historical and may be measured with error. Second, we may have inadequately controlled for all pre-1860 covariates that simultaneously affect slave proportion in 1860 and political attitudes today, which might result in a spurious relationship. To allay some of these concerns, we present additional robustness checks, matching analyses, and counterfactual comparisons between the North and the South in the Appendix. Here, we present results from an alternative specification that instruments for slave proportion in 1860 with county-level measures of the environmental suitability for growing cotton. We constructed these measures using data from the United Nations Food and Agriculture Organization (FAO).¹² Table 2 presents our instrumental

¹²These measures represent the maximum potential cotton yield based on soil, climate, and growing conditions. The estimates are based on climate averages from 1961 to 1990 and we average the “intermediate” and “high” levels of inputs, which refers to the effort required to extract the resource. We omit suitability for other crops, such as tobacco, because they have no relationship with slavery conditional on cotton suitability. While these measures use data from the contemporary period, we expect that most of the changes to the suitability between 1860 and 1960 to either be uniform shifts across the entire region due to worldwide climate change or be unrelated to local political attitudes.

variable (IV) estimates of the effects of proportion slave on the four outcome measures using a two-stage least squares (2SLS) model with state fixed effects, log of the county size, ruggedness of the terrain, water access, latitude and longitude, and their squared terms included as controls in both stages.¹³ Column (1) presents the first-stage relationship between cotton suitability and proportion slave. Columns (2) - (5) present the second stage estimates of the effect of proportion slave on the outcome measures. The results show second-stage estimates that are stronger than our baseline estimates, reported in Table 1.

For the IV approach to serve as a plausible identification strategy, cotton suitability must have an effect on contemporary attitudes exclusively through slavery, a strong assumption. Cotton suitability could, for example, determine how rural a county is today, which in turn could affect political attitudes. While the exclusion restriction is an untestable assumption, we assess its plausibility using a falsification test motivated by [Nunn and Wantchekon \(2011\)](#). We first estimate the reduced-form relationship between cotton suitability and contemporary beliefs both within and outside of the South—that is, mostly in the North. The legal absence of slavery in the North in this time period means that cotton suitability cannot affect political attitudes through slave prevalence. Any relationship between cotton suitability and political attitudes in the North would be a direct effect of cotton suitability on political attitudes. In [Appendix B.2](#), we present results from this analysis showing that, outside the South, the relationship between cotton suitability and political attitudes is either very small or in the opposite direction as in the South. This lends some additional credibility to our IV approach.

5 Explanations for Slavery’s Effect on Contemporary White Attitudes Based on Existing Theories of Political Behavior

We now turn to explaining the above findings. We first consider three possible explanations rooted in contemporary factors, as opposed to historical ones: racial threat, geographic sorting, and contemporary income inequality between blacks and whites, which could lead to statistical discrimination. (We examine two other explanations—rural/urban differences and the effects of Civil War destruction—in the [Appendix](#).) Af-

¹³To strengthen the internal validity of our design and minimize the potential for confounding, we omit counties with zero cotton suitability from the analyses. We use county-level data for the ANES measures to side-step the issue of clustering by county in the 2SLS model.

Table 3: Effects of slavery on white attitudes net the effect of the contemporary proportions of African Americans

	Prop. Democrat		Affirm. Action		Racial Resentment	
	(1)	(2)	(3)	(4)	(5)	(6)
Prop. Slave, Direct Effect	-0.150** (0.043)	-0.130* (0.057)	-0.153** (0.037)	-0.130** (0.042)	0.514** (0.140)	0.472* (0.184)
Prop. Black, 2000	0.172** (0.043)		0.062 (0.037)		-0.416** (0.140)	
State Fixed Effects	✓	✓	✓	✓	✓	✓
1860 Covariates	✓	✓	✓	✓	✓	✓
Bootstrapped SEs		✓		✓		✓
Model	WLS	Seq. g-est.	WLS	Seq. g-est.	WLS	Seq. g-est.
N	1,152	1,152	1,152	1,152	1,027	1,027
R ²	0.214	0.209	0.103	0.095	0.151	0.144

Note: †p < .1; *p < .05; **p < .01. Columns (1), (3), and (5) simply include proportion black in the year 2000 as an additional control to the baseline specification from Table 1. Columns (2), (4), and (6) use sequential g-estimation of Vansteelandt (2009).

ter finding no support for these contemporary factors, we turn in Section 6 to what we believe is the more likely explanation concerning the long-term historical persistence of political attitudes.

5.1 Racial Threat (Contemporary Black Concentrations)

As noted above, one plausible explanation for our results is that they are driven by contemporary black concentrations, which correlate highly with slavery. Indeed, the correlation between percent slave in 1860 and percent black in 2000 is 0.77. Since the local prevalence of slavery has produced high concentrations of blacks in the modern-day Black Belt, whites living in the Black Belt could be more conservative today simply due to modern racial threat.

To assess whether modern racial threat accounts for our findings, we check how much of our baseline results can be explained by contemporary black concentrations. We do so in two ways. First, we include the “mediator” (here, proportion black in 2000 as measured by the 2000 U.S. Census) as a covariate in the baseline specification, along with the treatment of interest (proportion slave in 1860). This analysis is shown in Table 3, columns (1), (3), and (5). The coefficient on proportion slave in 1860 remains significant and actually strengthens, suggesting that its direct effect does not operate through proportion black in 2000.

These estimates, however, could suffer from post-treatment bias (Rosenbaum, 1984). After all, the modern geographic distribution of blacks is a direct consequence of the prevalence of slavery (as noted by Key, 1949, for example). In addition, including the proportion black today as a mediator in a mediation approach would violate the key assumption of no intermediate confounders (Imai et al., 2011). We address these twin concerns by using a method developed in biostatistics by Vansteelandt (2009) and applied to other social science contexts by Acharya, Blackwell and Sen (2016). This method enables us to calculate the controlled direct effect of slavery, which is the effect of slavery on our outcomes if we were to fix the modern-day concentration of African Americans at a particular level (see Acharya, Blackwell and Sen, 2016, for a technical overview of the methodology). To implement the method, we use a two-stage estimator, called the sequential g-estimator, that estimates controlled direct effect of slavery when we have a set of covariates that satisfies the assumption that there exist no omitted variables for two relationships: one between the outcome and proportion slave in 1860 and the other between the outcome and proportion black in 2000 (Vansteelandt, 2009; Acharya, Blackwell and Sen, 2016).¹⁴ To estimate the direct effect of slavery, we first estimate the effect of contemporary black concentrations on whites' views today, controlling for all of our 1860s covariates and including the additional intermediate covariates in footnote 14.¹⁵ We then transform each of the outcome variables by subtracting the effect of black concentrations to create counterfactual estimates of the outcomes as if all counties had the same proportion black today.¹⁶ Finally, we estimate the effect of proportion slave on this transformed variable, which gives us the controlled direct effect of proportion slave on our measures of contemporary political attitudes.

Estimates from this analysis are reported in columns (2), (4), and (6) of Table 3.¹⁷ Compared to the baseline estimates of Table 1 and the potentially biased estimates in columns (1), (3), and (5), these results demonstrate that contemporary black shares

¹⁴That is, we want to control for intermediate factors that possibly led from slavery in 1860 to black concentrations today (and that are correlated with white attitudes today). To determine this set of covariates, we draw on the racial threat literature (see, for instance Giles and Buckner, 1993) and include log population in 2000, unemployment in 2000, percent of individuals with high school degrees in 1990, and log median income in 2000. These results assume no interaction between proportion slave and contemporary proportion black, but weakening this assumption does not change the findings (Acharya, Blackwell and Sen, 2016).

¹⁵Results from this first-stage can be found in Appendix Table A.12.

¹⁶That is, we take the coefficient on proportion black, multiply that by the actual share black in the county, and then subtract this product from each of the three white attitudinal outcome variables.

¹⁷To account for the added uncertainty of the two-step nature of sequential g-estimation, we report bootstrapped standard errors.

have little influence on slavery's effect on any of the outcomes. The direct effects of slave proportion are similar to those in Table 1 and are still highly significant. Moreover, once we account for slavery in 1860, contemporary black concentrations appear to have the opposite effect that racial threat theory would predict for Southern white attitudes. Finally, with the full controls from the first stage of the sequential g-estimator, the effect of proportion black today is no longer significant (Appendix Table A.12). Thus, we see no evidence that slavery's effects operate via contemporary black concentrations.

5.2 Geographic Sorting

The next possibility is that population sorting explains our results. For example, racially hostile whites from other parts of the South (or elsewhere) may have migrated to former slave counties during the last 150 years. Analogously, whites who hold more racially tolerant beliefs may have continually left former slaveholding areas.

We address this sorting hypothesis in several ways. First, for geographic sorting to explain our results, patterns of mobility into (and out of) the former slaveholding areas would have to differ from non-slaveholding areas. To investigate this possibility, we look into patterns of migration in a five-year snapshot from 1935-1940, drawing on the public use micro-sample (PUMS) of the 1940 U.S. Census (Ruggles et al., 2010). This year of the census is unique in that it provides the county in which a person resided in 1935 and in 1939. These data allow us to investigate if white migrants into or out of former slave areas during this time period were somehow distinct from other white migrants. If sorting plays an important role in our results, we would expect to see differences between migrants to/from high-slave areas versus low-slave areas. To test for differences among *out-migrants*, we adopted the following strategy: we ran a regression of individual characteristics on out-migration status for white respondents, the proportion of slaves in the respondent's 1935 county of residence, and the interaction between the two. We also included the 1860 covariates and state fixed effects for the 1935 counties. The interaction in this regression measures the degree to which differences between out-migrants and those who didn't migrate varies as a function of proportion slave. For *in-migration*, we take a similar approach but replace the characteristics of the 1935 county of residence with the characteristics of the 1939 county of residence.

Figure 3 depicts the results from this analysis, and shows the differences between white migrants and non-migrants across a number of characteristics.¹⁸ The figure de-

¹⁸See the IPUMS documentation at <https://usa.ipums.org/> for a complete description of these cen-

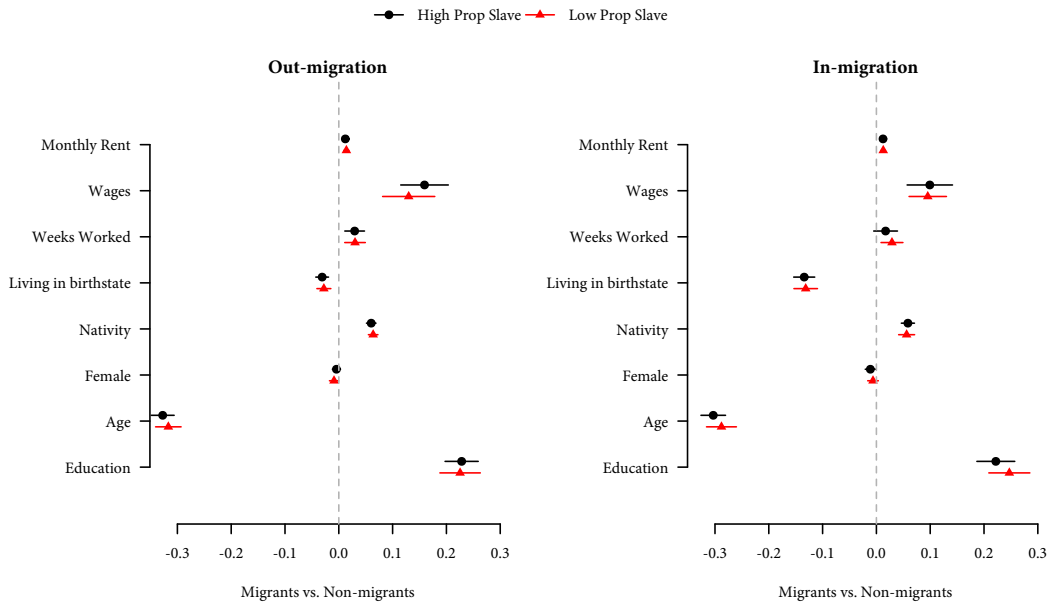


Figure 3: Characteristics of white out-migrants and in-migrants compared to white non-migrants for high-slave and low-slave counties, where migration took place between 1935 and 1939. In the left panel, each point is the estimated difference between non-migrants and out-migrants from high-slave areas (black dots) and between non-migrants and out-migrants from low-slave areas (red triangles), conditional on 1860 covariates of the individual's 1935 county of residence. The right panel is the same for in-migration, conditional on 1860 covariates of the individual's 1939 county of residence.

picts how these effects vary by proportion slave. For continuous outcomes, the effects are in terms of standard deviations; for binary outcomes, they are in terms of differences in proportions. With age, for example, this figure shows that both out-migrants and in-migrants are significantly younger compared to those who remained in their counties. However, the key point is that this pattern is the *same* for low and high-slave counties, as evidenced by the overlapping estimates. Thus, migrants into or out of high-slave counties are not any younger than migrants from or to other counties. This pattern holds more generally: migrants in the 1935-1940 period are distinct from non-migrants, but those differences are fairly constant across proportion slave in the county. Indeed, across all of these characteristics, there are no significant interactions between migrant status and proportion slave, meaning that migrants from/to high-slave and low-slave areas appear very similar to each other. These results are specific to a specific five-year period in American history for which we have data, but they do provide sus measures.

suggestive evidence that the differences we are seeing are not due to geographic sorting alone.

Second, we also provide additional evidence in Appendix D, in which we analyze contemporary mobility data from the 2000 U.S. Census. These analyses show mobility mostly from formerly low-slave to low-slave areas or from formerly high-slave to high-slave areas (as opposed to mobility between former low-slave and former high-slave areas, or vice versa). Taken together, the evidence suggests that geographic sorting is unlikely to be the exclusive explanation behind our results.

5.3 Inequality and Statistical Discrimination

The final alternative explanation that we consider here is that our regional patterns are the outcome of the contemporary economic landscape—in particular, contemporary economic inequality between blacks and whites. O’Connell (2012), for example, finds that slavery in 1860 predicts black-white income inequality today (results echoed by Nunn, 2008; Lagerlöf, 2005). Her results suggest that our findings might be explained by a theory of statistical discrimination (Becker, 2010). In addition, because poverty correlates with other traits (such as, for example, higher instances of crime) then we might expect this type of statistical discrimination to be even more pronounced. Consistent with O’Connell (2012), we show in Appendix Table A.14 that slavery has an effect on black-white income inequality measured using both 1940 wages and 2014 median incomes, though the effects have attenuated substantially since 1940. Could this income inequality between blacks and whites in the Southern Black Belt be driving our findings on whites’ political attitudes, for example through statistical discrimination (discrimination against blacks because they are, on average, poorer than whites)?

We address this question with two analyses, both of which cast doubt that income inequality and statistical discrimination explain our results. First, we show that local income inequality is a weak predictor of our main outcome variables, and, if anything, the effects are in the opposite direction as statistical discrimination theory would predict. These are shown in Table 4, which shows that inequality in 2014 is weakly related to political outcomes and when it is related, the relationships are in the opposite direction of what would explain our results. In other words, counties with higher black-white inequality are places where whites are more likely to be Democrats, more likely to support affirmative action, and are about as racially resentful as whites elsewhere. Furthermore, the table also shows that including contemporary black-white income inequality as a

Table 4: Effect of slavery versus the effect of inequality

	Proportion Democrat		Support for Affirm. Action		Racial Resentment	
	(1)	(2)	(3)	(4)	(5)	(6)
log White-Black Income Ratio, 2014	0.040* (0.017)	0.056** (0.019)	0.027 [†] (0.015)	0.041* (0.016)	-0.088 (0.057)	-0.092 (0.062)
Prop. Slave, 1860		-0.126** (0.044)		-0.160** (0.038)		0.475** (0.146)
State Fixed Effects	✓	✓	✓	✓	✓	✓
1860 Covariates	✓	✓	✓	✓	✓	✓
N	1,140	994	1,140	994	1,017	894
R ²	0.180	0.227	0.075	0.125	0.069	0.159

Note: [†]p < .1; *p < .05; **p < .01. Inequality here is measured by the log of the ratio of white to black median incomes within counties in 2014. Data comes from the American Community Survey, 2009-2014.

covariate in our baseline specification does not substantially change our estimates.¹⁹

Second, we check an observable implication of the inequality account. If income inequality drives slavery's effects, then we would expect the marginal effect of slavery on attitudes to vary in the income level of the respondent. For example, higher income respondents might be more (or less) discriminatory on the basis of black-white income inequality. In Appendix Table A.15, we show that there is no interaction between respondent income and proportion slave. Taken together, these analyses provide evidence against income inequality and economic conditions being the main mechanism behind our findings.

6 An Explanation for Slavery's Effect on Contemporary White Attitudes Based on the Historical Persistence of Political Attitudes

If the above explanations, largely rooted in contemporary factors or population sorting, do not explain our findings, then what does? In this Section, we lay out the theory of historical persistence in political attitudes. In our context, this is the idea that slavery and its collapse reinforced regional differences in political attitudes, and these differences have persisted over time.

¹⁹The coefficients on proportion slave in these models, however, should be interpreted with caution due to the potential for post-treatment bias.

The prevalence of slavery, coupled with the shock of its removal, created strong incentives for Black Belt whites to try to preserve both their political and economic power by promoting racially targeted violence, anti-black norms, and, to the extent legally possible, racist institutions. These reinforced racial and political beliefs about black subjugation within the Southern Black Belts, which, via institutional path dependence and intergenerational socialization, have persisted to the present day. While there is no question that anti-black attitudes were rampant throughout the South before the Civil War, we show in this Section that even for counties that were politically similar before the War, differences in partisan voting became more pronounced in the late 19th and early 20th centuries. Thus, although racism was prevalent across the high-slave and low-slave South in the antebellum period, these areas diverge greatly in terms of both institutionalized and socially enforced racism around the time of the Civil War, specifically the period of Redemption that followed Reconstruction (Foner, 2011; Logan, 1954).²⁰

Below, we draw on work in both history and in economics to provide some qualitative background to the way in which Southern white elites had incentives to foster and sustain anti-black suppression in the former bastions of slavery. We then provide evidence for our mechanism by presenting evidence on (1) the burgeoning political importance of race in the postbellum period, (2) postbellum racial violence (including lynchings) and economic oppression, (3) the weakening of economic incentives for racial hostility that took place as a result of the movement toward agricultural mechanization beginning in the 1930s, and (4) evidence for inter-generational transmission of racial attitudes.

6.1 Political and Economic Incentives for Black Repression in the Postbellum Years

Why would racial attitudes in the Black Belt persist after the Civil War, even as other regions of the country gradually changed their views on race? Our explanation for this postbellum divergence lies in the fact that, after emancipation, Black Belt white elites faced two interrelated threats. The first was political. In areas where blacks outnumbered whites (in some cases 9 to 1), the abrupt enfranchisement of blacks threatened

²⁰The more general idea behind this theory is that when an entrenched social and economic institution like slavery is abruptly and forcibly abolished, previously powerful groups (ex-slave-owning white elite) seek to establish other local and informal institutions that serve a similar purpose to that of the previous, forcibly abolished formal institution (slavery) (e.g., Acemoglu et al., 2011).

white control over local politics (Du Bois, 1935; Key, 1949; Kousser, 1974). This gave whites in former slaveholding counties an incentive to promote an environment of violence and intimidation against the new freedmen, with the purpose of election fraud and disenfranchisement (Du Bois, 1935; Kousser, 1974). The second threat to white elites was economic. The emancipation of slaves after the Civil War was a major shock to the Southern economy: blacks now had to be paid wages (Higgs, 1977). Furthermore, emancipation brought blacks some freedom over the amount of labor they supplied, and many ex-slaves chose to work for themselves rather than for the white ruling class (Ransom and Sutch, 2001*b*). This both reduced the labor supply and increased labor costs sharply, threatening the Southern plantation economy (Ransom and Sutch, 2001*b*; Alston and Ferrie, 1993).²¹ Whites therefore had an incentive to establish not just new forms of labor coercion that could replace slavery but also new political restrictions that would help protect white hegemony.

Since black populations and large-scale agricultural production were greatest in former slaveholding counties, it was in these counties that Southern elites exerted greater efforts toward repression (Kousser, 1974). These repressive techniques are well documented in the economics and history literatures (Alston and Ferrie, 1993; Blackmon, 2008; Lichtenstein, 1996; Wiener, 1978). For example, Wiener (1978, p. 62) describes how “planters used [Ku Klux] Klan terror to keep blacks from leaving the plantation regions, to get them to work, and keep them at work, in the cotton field.” Also well documented is the fact that poor whites were complicit with the landowning elite and would engage in and support violent acts towards blacks, even though such violence could presumably also lower white wages (Du Bois, 1935; Blackmon, 2008; Roithmayr, 2010). Note that these forces are distinct from the modern interpretation of racial threat, under which whites would find high concentrations of black populations undesirable. In our explanation, the Southern white population sought to control, police, and dominate black labor resources. Blacks were not threatening because of their presence (as racial threat would imply); they were threatening because they were an important provider of labor and, in the post-emancipation environment, they could leave.

Three additional factors sustained and nurtured anti-black attitudes in the former slaveholding counties, both within the community and over time. The first is the continued and perhaps increased use of social norms that directly or indirectly put blacks

²¹Some of these concerns were mitigated by the sharecropping system that became pervasive in the post-bellum period. This aligned the incentives of the planters and the laborers to some extent (Ransom and Sutch, 2001*b*, p.88-89), but these arrangements did not reduce the incentives of planters to further bolster their position in the labor market.

in an inferior position relative to whites (Logan, 1954; Du Bois, 1935).²² The second is the expansion of these attitudes across local white communities, which could operate in a manner consistent with existing theories of the diffusion of political opinions from elites to the general public (Zaller, 1992). The last is the intergenerational transfer of political and racial attitudes, in a manner consistent with theories of intergenerational socialization both in economics and cultural anthropology (Boyd and Richerson, 1988; Bisin and Verdier, 2000) as well as in political science (Campbell et al., 1980; Jennings and Niemi, 1968).²³ In our context, attitudes could be passed down from one generation to the next through both cultural and institutional channels, for example via institutions such as Jim Crow or socially enforced segregation and racially motivated violence. In line with the literature, such transmission would necessarily be imperfect, suggesting that there is some decay in these geographically-based relationships over time. This leads to one important empirical prediction: the effects of slavery should be smaller in areas where the incentives for anti-black attitudes faded earlier. Below, we provide evidence for this prediction.

Although our discussion focuses on the postbellum South, our argument is not that racial hostility was nonexistent before the Civil War. Instead, the incentives that gained strength at the time of emancipation likely exacerbated the political differences between former slaveholding and non-slaveholding areas. Racially hostile attitudes, in other words, may have dissipated more quickly in areas that were non-slaveholding and were less reliant on the provision of inexpensive black labor. The South was surely racist before the Civil War, but for this to explain our results, the strength of the racism would have to align with the density of slavery. Our previous analyses provide some limited evidence against this idea: in results presented above, for example, our effects are robust to controlling for the antebellum presence of freed blacks in the county, which may be possible indicator of antebellum racial attitudes. They are also robust to the inclusion of a host of antebellum factors capturing economic and political differences. Below, we

²²The political and economic incentives for racial violence and oppression is likely to have produced or reinforced racially hostile attitudes among whites through psychological and other channels. For example, whites might have developed or maintained racially hostile attitudes to minimize the “cognitive dissonance” associated with racially-targeted violence towards blacks (Acharya, Blackwell and Sen, 2015). Theories in social psychology, beginning with the work of Festinger (1957), would suggest that engaging in violence could produce, or sustain, hostile attitudes among members of the perpetrating group towards the victim group, if individuals from the perpetrating group seek to minimize such dissonance.

²³We note that under our mechanism, it is socio-political *attitudes*, rather than *partisanship*, that are passed down from parent to children. This makes our mechanism consistent with the partisan realignment that took place in the 1960s, given the assumption that partisanship depends at least in some part on attitudes.

focus more closely on these ideas.

6.2 Timing of the Political Divergence and the Importance of the Reconstruction Period

What additional evidence do we have that racial attitudes became politically salient in the postbellum period? To shed light on this question, we examine the relationship between proportion slave in 1860 and a long-standing historical measure of political attitudes, presidential vote shares. Until realignment in the middle of the 20th century, the Democratic Party was the racially conservative party, while the Republican Party was the racially progressive party (Black and Black, 1987). In the period before the Civil War, the Second Party system of Democrats and Whigs attempted to remove slavery as a national issue, but slaveholding still drove political differences on issues such as internal improvements and tariffs, with elite slaveholders more likely to support Whig candidates (Holt, 1999, pp. 115–121). Do these differences drive the relationship between slavery and modern attitudes? Or do post-War events drive the results we see above?

We address these questions by examining the effect of slavery on county-level Democratic vote share of the presidential vote between 1844 and 1968.²⁴ To analyze the time trend, we calculate the effect of a 25 percentage-point increase in 1860 slavery on the presidential vote in each year. Thus, each point in Figure 4 represents a (scaled) point estimate from a regression of county-level Democratic vote share on county proportion slave in 1860, using the same cotton-suitability IV design as Table 2.²⁵ As the figure shows, there is little difference between slave and non-slave areas before the Civil War, with the exception of 1856 where, if anything, high slave areas are more likely to vote for the more moderate candidate on slavery, former president Millard Fillmore (compared to the relatively more pro-slavery Democratic candidate James Buchanan). Thus, at least in terms of national party politics, differences in white views appear to organize around the density of slavery more strongly after the Civil War.²⁶

²⁴Since election outcomes are not disaggregated by voter race, these data also include black voters. Blacks voted in large numbers following emancipation but were systematically disenfranchised between in the late 19th and early 20th centuries. Large scale re-enfranchisement did not occur until the Voting Rights Act of 1965, which is why we stop the analysis of such vote shares at that time.

²⁵We use the original 1860 county boundaries for the measurement of slavery for outcomes until 1924, after which we use the data interpolated to modern boundaries. The changes in county boundaries between 1924 and today in the sample states is minimal.

²⁶We obtain similar results for congressional elections and using the WLS specification.

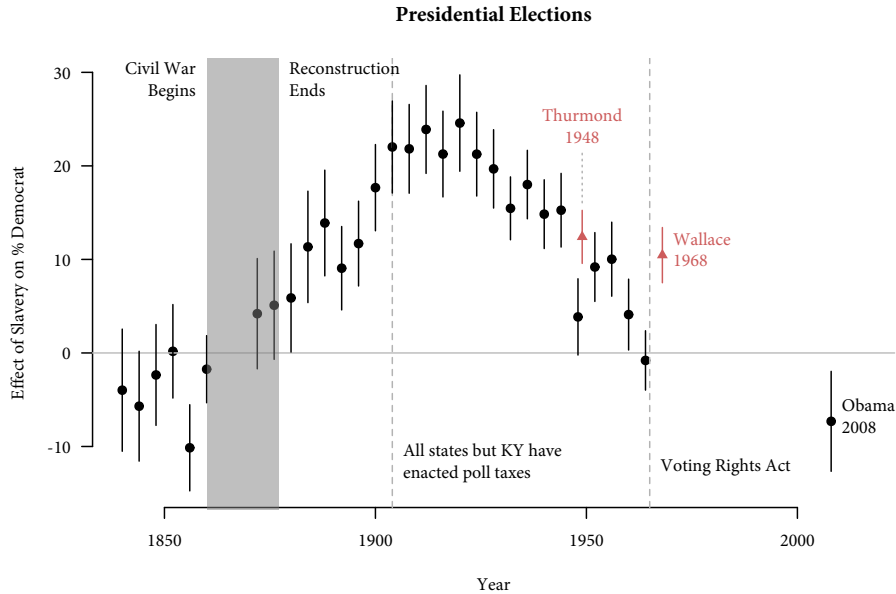


Figure 4: Effect of proportion slave on vote for Democratic presidential candidate in the South over time. Each point is the effect of a 25 percentage-point increase in proportion slave from separate IV models of county-level Democratic share of the presidential vote on proportion slave. Results for Obama in 2008 are from White respondents in the CCES.

Instead, as Figure 4 indicates, the data show that regional differences emerged after the end of Reconstruction. As white elites begin to restrict the vote of African Americans in the late 19th century, the effect of slavery increases, reaching its peak around the time that most of the states finalized the enactment of poll taxes and literacy tests to almost fully disenfranchise blacks (Kousser, 1974). By around 1950, the effect of slavery weakens, in part due to some small additions of African Americans to the registered voter pool,²⁷ and also to the move of national Democratic candidates toward a civil rights platform. The effect sizes from the 1950s and 1960s are roughly similar in magnitude, but in the opposite direction, to the effect of slavery on the Obama white vote in 2008, estimated from CCES white respondents. Throughout, the difference in voting behavior between formerly large slaveholding counties and other counties is large and statistically significant. These findings also hold if we use the WLS specification

²⁷The percentage of the Black Voting-Age Population expanded from 3% to 18% in Georgia, and from 0.8% to 13% in South Carolina (Mickey, 2015).

and condition on our pre-treatment covariates, and they hold even when we control for county-level vote for Democrat James Buchanan in the 1856 presidential election. These findings suggest that large and significant differences in partisanship emerge in the postbellum period even among counties that were politically and economically similar in the antebellum period. Furthermore, they reinforce the historical account of [Foner \(2011, pp. 11–18\)](#), who describes how the Civil War amplified divisions and resentments within the white Southern population, concluding that “the [Civil War] redrew the economic and political map of the white South” (p. 17).

However, regarding antebellum differences, we note that presidential election results in the prewar period could be masking important intraregional differences, especially since the Second Party system attempted to remove slavery as an issue in national politics. For a stronger test of the effect of slavery on antebellum politics, we turn to two state-level elections that centered on slavery and sectional issues. Specifically, after the Compromise of 1850, many states in the South were divided over the deal’s fairness and whether to extend popular sovereignty to the territories. The deal, which admitted California as a free state, proved so unpopular that several gubernatorial elections in 1851 pitted states’ rights extremists, who openly discussed secession, against more moderate Unionist candidates, who supported the compromise ([Cole, 1914, pp. 181-189](#)). In Georgia, for example, Constitutional Union candidate Howell Cobb ran against former governor and Southern Rights candidate Charles McDonald, and, in Mississippi, moderate Senator Henry Foote ran against former Senator (and future Confederate president) Jefferson Davis. These elections provided the (white) electorate a clear choice in the national fight over slavery and could potentially reveal antebellum regional differences on the issue. [Figure 5](#) presents the relationship between slavery in 1850 and county-level vote-shares for the moderate candidates. In both states, perhaps surprisingly, there is little evidence of a strong relationship between slavery and vote choice, even in an election that focused so heavily on the issue. Low-slave areas were not actively anti-slavery and did not appear to actively resist secession in the early 1850s. Furthermore, the high-slave areas were at least as willing to compromise on the issue of slavery, suggesting that the contemporary political geography of slavery probably does not trace its origins to this time period.

6.3 Lynchings and Other Forms of Suppression

Our incentives-based mechanism posits that violence was used not only to disenfranchise blacks, but also to suppress their mobility and wages—a particularly strong in-

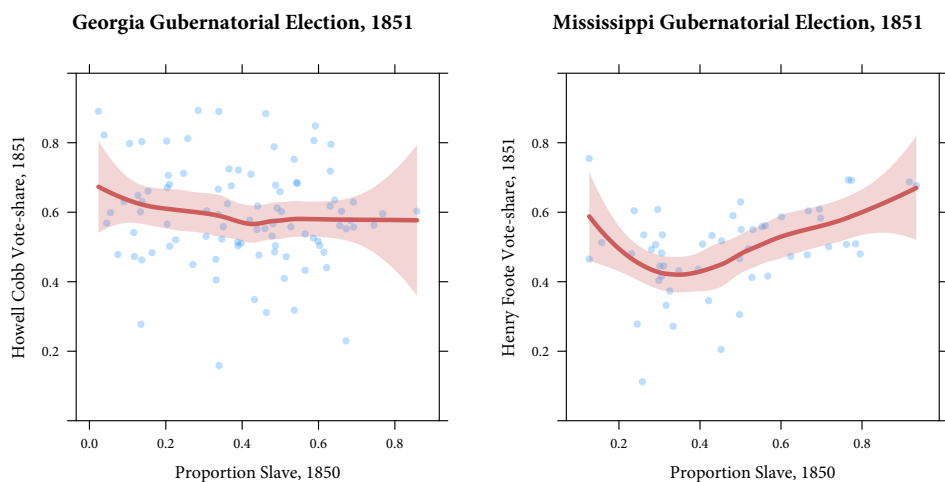


Figure 5: Within-state relationship between proportion slave in 1850 in a county and percentage voting for Unionist candidates in 1851 in that county. Both Howell Cobb and Henry Foote were running against strong States' Rights supporters in the aftermath of the Compromise of 1850. Higher values on each y -axis indicate more moderate voting with regard to slavery and secession.

centive in the postwar, post-emancipation period. If the mechanism is at play, then we would expect to see greater racist violence in former slaveholding counties in this time period. While we do not have measures of all forms of violent racism in the post-Reconstruction era, we do have county-level measures of one extreme form of racial violence: lynchings (Beck and Tolnay, 2004).²⁸

In column (1) of Table 5, we confirm the hypothesis that the number of black lynchings between 1882 and 1930 per 100,000 1920 residents is greater in counties that had high slave proportions in 1860, conditional on state-level fixed effects and our 1860 covariates. The effect of slavery in 1860 and lynchings in this period is large and significant: a 10 percentage-point increase in slave proportion is associated with a 1.89 increase in lynchings per 100,000 residents. This result is in line with our incentive-based theory: there is more racial violence in areas previously more reliant on slave labor. Furthermore, if our explanation is correct, then black farmers should be worse off in former slave areas due to this greater local violence. Appendix Table A.17 draws on data from the the 1925 Agricultural Census (Haines, 2010) and shows that, in comparison to white farmers, black farmers in former high-slave areas were significantly worse off than those in other areas of the South. They were more likely to be under

²⁸These data include all states in our analysis except Texas, Virginia, and West Virginia.

Table 5: Effect of slavery on postbellum violence and effect modification by mechanization.

	Lynchings (1)	Prop. Democrat (2)	Affirm. Action (3)	Racial Resentment (4)
Prop. Slave, 1860	17.656** (5.661)	-0.183** (0.049)	-0.219** (0.042)	0.703** (0.161)
Tractors Change, 1930-1940		-0.417† (0.243)	-0.629** (0.207)	2.290** (0.791)
Prop Slave × Tractors Change		2.298** (0.868)	2.226** (0.741)	-7.762** (2.876)
State-Year Fixed Effects	✓	✓	✓	✓
1860 Covariates	✓	✓	✓	✓
N	910	1,145	1,145	1,020
R ²	0.350	0.208	0.104	0.152
Specification	WLS-county area	WLS-sample size	WLS-sample size	WLS-sample size

Note: †p < .1; *p < .05; **p < .01. The first column is WLS with the total county area as weights (and where county area is omitted from the 1860 covariates, though this has no effect on our analysis). The remaining columns are WLS with within-county sample size as weights. Lynchings are black lynchings between 1882 and 1930 per 100,000 1920 residents (similar results hold using average population size between these dates). Tractors change is the change in tractors per 100,000 acres of land between 1930 and 1940. Tractors in 1930 is the number of tractors per 100,000 acres of land in 1930.

tenancy agreements and less likely to own their own farm.

6.4 Mechanization of Southern Agriculture

After slaves were freed, Southern elites used various tactics to gain an advantage in the market for black labor, including racial violence, convict leasing, anti-vagrancy laws, and anti-enticement laws ([Ransom and Sutch, 2001a](#)). A clear implication of our incentives-based mechanism is that once the demand for black labor drops due to exogenous technological development, the incentives for whites to interfere in the labor market with such tactics should lessen, and thus the effects of slavery on attitudes should also diminish. This implication is testable given that much of the Southern economy was agricultural, and its main cash crop, cotton, was heavily labor intensive until about the 1930s when Southern agriculture started to mechanize and tractors began to replace labor.

To test this, we examine whether counties that mechanized earlier are indeed those where the effects of slavery wanes more quickly. Following [Hornbeck and Naidu \(2014\)](#), we use the number of farming tractors as a proxy for mechanization.²⁹ We interact the

²⁹Tractors were an early form of mechanization for the cotton industry, intended to replace mule-

proportion of slaves in 1860 with the change in the number of tractors per 100,000 acres of agricultural land in the county between 1930 and 1940, which we collect from the 1930 and 1940 Agricultural Censuses (Haines, 2010). We then estimate the effect of this interaction on our three outcome measures. As Table 5 shows, the effects of slavery are weaker for counties where mechanization grew between 1930 and 1940. For example, where mechanization did not grow between 1930 and 1940, a 10 percentage-point increase in proportion slave leads to a 1.6 percentage-point drop in the percent of whites who identify as Democrat today (95% confidence interval: $[-2.5, -0.8]$). Where mechanization grew rapidly, with 0.07 more tractors per 100,000 acres (95th percentile), the same change in proportion slave leads to only a 0.2 percentage-point decrease in the percent Democrat (95% confidence interval: $[-1.1, 0.07]$).

We note two potential concerns with this test. First, the results could be consistent with a racial threat explanation—early mechanization led to decreases in the black population in these areas (as shown by Hornbeck and Naidu, 2014). In Appendix Table A.16, however, we replicate the analysis using the declines in proportion black from 1920 to 1940 and 1970 and find that areas with larger declines have, if anything, larger effects associated with slavery. Thus, it seems unlikely that diminishing racial threat is driving the attenuating effects we see above. Second, it could be that more racially tolerant counties chose to mechanize early in order to rid themselves of the incentives for racial exploitation. However, as shown in Appendix Table A.18, there is no relationship between growth in tractors and either racial violence or inequality in wages between blacks and whites. This casts doubt that tractors or their growth are indicators of racial attitudes. Moreover, as Hornbeck and Naidu (2014) argue, many of the counties that mechanized early were those affected by an exogenous shock, that of the Mississippi floods of 1927, and are thus *ex ante* similar to counties that mechanized later.

6.5 Intergenerational Transmission of Beliefs

Our last empirical analysis concerns the intergenerational transfer of political attitudes—or, how attitudes on politics and race have been passed down over time. One possibility is that attitudes are shaped contemporaneously by local institutions, for example schools and churches, which have themselves persisted. Another possibility is that these attitudes have been passed down from parents to children, perhaps within a broader

drawn plows (Wiener, 1978). Full-blown mechanization of cotton production via the cotton picker would not become widespread until after 1940.

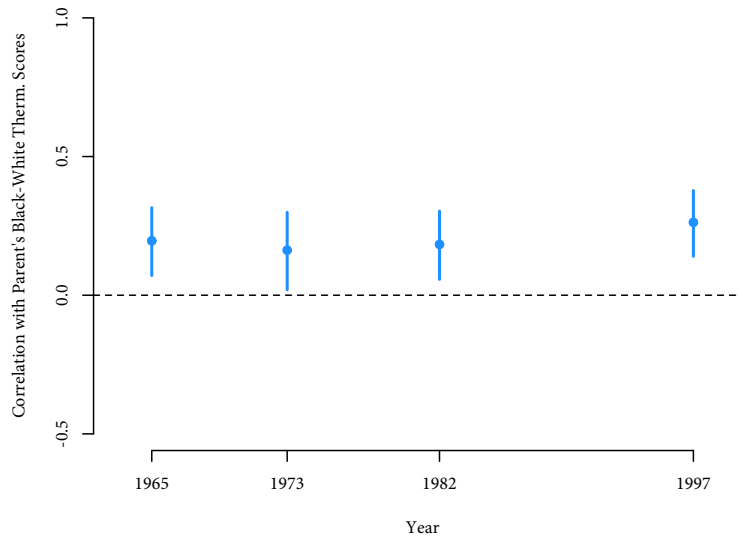


Figure 6: Correlation coefficients between parent's white racial preference in 1965 and their children's white racial preference measured at four time periods: 1965, 1973, 1982, and 1997. White racial preference here is measured as the difference between the respondent's thermometer scores for white and black people.

context of institutions, but also through intergenerational socialization. Both transmission mechanisms are likely at play, and finding correlated attitudes between parents and children would be consistent with research in political science demonstrating the importance of parents' partisanship in shaping children's partisanship (e.g. [Campbell et al., 1980](#); [Jennings and Niemi, 1968](#)).

To explore the inheritance of racial attitudes from previous generations, we rely on the Youth-Parent Socialization Panel Survey, which measured the racial attitudes of a national probability sample of high school senior students in 1965 along with their parents ([Jennings et al., 2005](#); [Elliot, 2007](#)). This study followed up with the students and interviewed them in 1973, 1982, and 1997, with response rates over 80% in each survey waves. We focus on the subset of this sample who lived in the South in 1965 ($n = 241$), which is a good proxy for students who were raised in the South.³⁰ If at-

³⁰See [Jennings, Stoker and Bowers \(2009\)](#) for a fuller analysis of the national results from this panel survey.

Table 6: Effects of slavery for those born after the Voting Rights Act.

	Prop Democrat logistic (1)	Affirm. Action logistic (2)	Racial Resentment OLS (3)
Prop. Slave, 1860	-1.113** (0.365)	-0.840* (0.358)	0.672* (0.265)
State-Year Fixed Effects	✓	✓	✓
1860 Covariates	✓	✓	✓
N	9,551	9,528	3,350
R ²			0.059

Note: †p < .1; *p < .05; **p < .01. All models at the individual-level with standard errors clustered on county and weighted by CCES survey weights.

titudes are being passed down from generation to generation, then we would expect that the racial attitudes of parents in 1965 should be correlated with the views of their children in each wave of the survey. We measure racial views in these surveys as the difference between the thermometer score for whites and the thermometer score for blacks, which is available for both parents and children in each wave. Figure 6 shows these correlations for each wave and demonstrates how stable this relationship is over time. What these correlations show is that children with racially conservative parents in 1965 are more likely to be racially conservative themselves at least through age 50, which is evidence of intergenerational socialization.

A final concern is that our findings regarding contemporary attitudes may be driven not by intergenerational transmission, but by the direct experiences and attitudes of older individuals in our CCES sample. That is, we are detecting an effect only among older whites, which would suggest no or limited intergenerational transfer of attitudes and instead an exposure to the tail end of Jim Crow. To test this, we estimate the effect among a subgroup who were more likely to receive such attitudes only from their parents: whites born after the Civil Rights Act of 1964 and the Voting Rights Act of 1965. While segregation and race-related oppression extended well past this time, both pieces of legislation have been acknowledged as influential in strongly reducing segregation in the South and increasing black enfranchisement (Rosenberg, 2008). In Table 6, we show the effect of slavery on these younger whites is just as strong and statistically significant as it is for older whites. Together with our results on racial threat and income-based discrimination (in Section 5), these findings provide some evidence that the parent-to-child transmission mechanism is an important component of how

slavery affects attitudes. As such, this evidence provides some support that political *culture*, rather than exclusively institutions, plays an important role in explaining the persistent effects of American slavery.

7 Conclusion

In this paper, we have shown that an institution that was formally abolished 150 years ago still has effects on attitudes today. Specifically, we show that American slavery has had a direct impact on Southern whites' (1) partisan identification, (2) attitudes on affirmative action, (3) levels of racial resentment, and (4) attitudes toward blacks. We further showed that our findings are robust to instrumenting for cotton suitability, lending credibility that our estimates are causal. Our findings are robust to accounting for a wide variety of factors that could plausibly affect both the share of the population that was enslaved and also contemporary political attitudes. In addition, we ruled out several contemporary-based explanations. Specifically, we ruled out the theory of "racial threat," or the idea that contemporary shares of the black (minority) population are what drive white (majority) group racial attitudes. When we took into account contemporary shares of the black population using appropriate methods, we found that slavery continues to have a separate direct effect. We also provided suggestive evidence that our results are not due exclusively to geographic mobility over the course of the 20th century, nor are they due to contemporary income inequality between blacks and whites.

Our results instead suggest a separate causal channel attributable to the historical persistence of regional variation in attitudes that was amplified by post-Civil War events. The years during and after the Reconstruction period saw whites coordinating to provide an informal social infrastructure (and to the extent legally permissible an institutional one as well) to maintain as much as possible the economic and political power previously guaranteed to them under slavery. As affirmative support, we showed that greater prevalence of slavery predicts more conservative (for many years more Democratic) presidential vote shares, higher rates of radical violence, and decreased wealth concentrated in black farms in the decades after Reconstruction. We also showed that the long-term effects of slavery are smaller in areas of the U.S. South that were quick to mechanize in the early to mid-20th century. Finally, we also offered evidence that parent-to-child transmission could be an important mechanism by which attitudes have been passed down over time. However, we do not rule out that Southern

institutions may have also played an important role.

Our research has substantial implications for our understanding and study of politics. Much work within political science, and on public opinion specifically, focuses on contemporary respondent factors, such as income, age, gender, education, and so on in trying to explain contemporary public opinion. However, this style of research overlooks historical and culturally-rooted explanations for the formation of public opinion. Our findings here suggest that historical institutions like slavery are significant in shaping American culture and politics, even if they no longer exist. In light of this, the field of political behavior could benefit from exploring other potential relationships between historical forces and contemporary attitudes. As Key (1949) himself observed, social and historical forces have “an impact on political habit whose influence has not worn away even yet.” This might be the case not just for development of political attitudes in the U.S. South, but also in other arenas within American politics and elsewhere in the world.

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A Supplemental Information: Data details

Areal interpolation. We use areal weighting to interpolate data from the 1860 U.S. Census onto modern county boundaries (Saporito et al., 2007). While population-weighted interpolation generally produces better estimates of the underlying rates of various sub-populations, we use areal weighting because we are interested in interpolating both proportions (enslaved, free, etc) and levels (total population, improve farming acreage, etc), and areal interpolation allows us to use a consistent set of interpolation weights across these disparate measures. O’Connell (2012) and Reece and O’Connell (2015) use a population-weighting approach to estimate proportion slave in modern county boundaries and our measure correlates with this measure at $r = 0.986$. For historical county boundaries, we relied on the Newberry Atlas of Historical County Boundaries (Siczewicz, 2011).

The areal weighting scheme works as follows. For a given census year, we create a $n_s \times n_t$ matrix A , where n_s is the number of source (1860) counties and n_t is the number of target (2000) counties. A is a row-normalized matrix, where each entry a_{ij} is the proportion of the area of source county i that is contained in target county j . We follow O’Connell (2012) and set $a_{ij} = 1$ and $a_{ij'} = 0$ if more than 95% of an 1860 county is contained in a single 2000 county. Let y_t be the vector of target values that we are trying to estimate and y_s be the observed source vector of values. Then, we construct areal weighted estimated by $y_t = A'y_s$. Essentially, this distributed the population in each 1860 county is distributed to 2000 counties based on how much of the 1860 county is contained in the 2000 county. Areal weighting assumes that distributions are evenly spread throughout the source counties, which may be false, and many methods exist to incorporate additional information such as roads to correct for these issues. Unfortunately, there is little information about the distribution of individuals or farms within counties in the antebellum period that could be reliably used across the entire country.

Antebellum transportation. We rely on the work of Atack (2013) and the resulting GIS shapefiles to identify navigable waterways (canals and steamboat-navigable rivers) and railroads as they existed in the 19th century (Atack, 2015a,b,c). To measure the availability of water-based transportation in modern counties, we detected if a canal or steamboat-navigable river crossed the boundaries of the county or if that county was on the coast of an ocean, the Gulf of Mexico, or one of the Great Lakes. A county meeting any of these criteria was measured as a 1, indicating it had water-based transportation available to it in 1860. A county was counted as having rail-based transportation if a railroad crossed through its territory.

Table A.1: Summary statistics for the county-level variables.

Statistic	N	Mean	St. Dev.	Min	Max
<i>Outcomes</i>					
Prop. Democrat	1,329	0.314	0.256	0.000	1.000
Support for Affirmative Action	1,329	0.222	0.227	0.000	1.000
Racial Resentment	1,168	4.116	0.701	1.000	5.000
Lynchings per 100,000 1920 Residents, 1882-1930	1,263	8.942	17.913	0.000	166.667
<i>Geographic Variables</i>					
County Area, 2000	1,326	0.151	0.090	0.001	1.497
Ruggedness	1,244	43.377	48.308	2.106	334.972
Latitude, 2000	1,326	34.528	3.068	24.850	40.521
Longitude, 2000	1,326	-87.297	6.675	-106.235	-75.685
<i>1860 Variables</i>					
Prop. Slave, 1860	1,242	0.284	0.214	0.000	0.924
Gini Coefficient for Land Holdings, 1860	1,222	0.481	0.084	0.000	0.789
Prop. Small Farms (< 50 Acres), 1860	1,226	0.466	0.215	0.019	1.000
Total Population, 1860	1,257	8,695.350	9,394.037	0.000	174,491.000
Farm Value per Capita, 1860	1,226	36.473	25.781	4.809	226.300
Prop. Free Black, 1860	1,242	0.010	0.022	0.000	0.246
Rail Access, 1860	1,326	0.245	0.430	0	1
Water Acces, 1860	1,326	0.480	0.500	0	1
<i>Other Variables</i>					
Cotton Suitability	1,323	0.428	0.191	0.000	0.921
Prop. Black, 2000	1,327	0.158	0.171	0.000	0.846
Tractor Growth, 1930-1940	1,295	0.020	0.032	-0.030	0.252

Note: Restricted to counties in the sample states (see text).

Table A.2: Summary statistics for individual-level data.

Statistic	N	Mean	St. Dev.	Min	Max
<i>CCES, 2006, 2008, 2009, 2010, 2011</i>					
Democratic Identification	40,154	0.341	0.474	0	1
Support for Affirmative Action	40,087	0.231	0.421	0	1
Racial Resentment	16,890	4.002	1.137	1.000	5.000
Age	40,154	52.083	15.380	18	96
Religious Importance	40,104	0.725	0.446	0	1
Gender (Female = 1, Male = 0)	40,154	0.524	0.499	0	1
Household Income (Bracket Midpoint)	35,584	62,074.810	42,603.620	5,000	175,000
<i>ANES, 1984-1998</i>					
Black Therm. Score	2,831	62.548	21.622	0	100
White Therm. Score	1,819	74.916	18.780	0	100
White-Black Therm. Difference	1,784	12.854	22.787	-50	100

Note: Restricted to self-identified whites living in the sample states (see text).

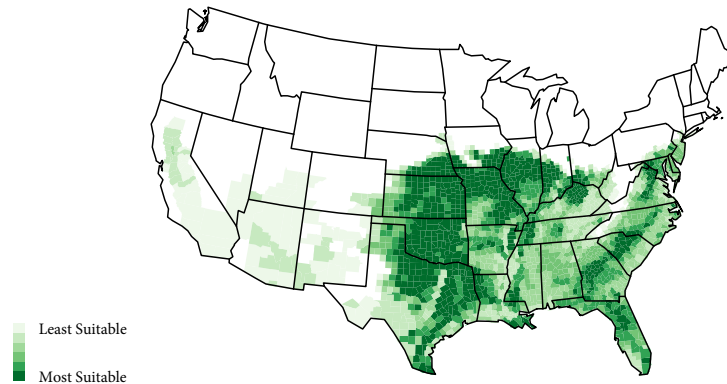


Figure A.1: Cotton suitability as evaluated by the U.N. Food and Agriculture Organization (FAO).

B Supplemental Information: Additional Analyses

B.1 Individual-Level Analyses

See Tables A.3 and A.4.

B.2 Falsification Test of the Instrumental Variable Exclusion Restriction

Figure A.1 presents the suitability of growing cotton in various areas of the country; this map shows that several non-slave areas of the country were suitable for cotton, including parts of the midwest (IL, IN, IA, and NE) and southwest (CA, NM, AZ, and OK).

We present the results of this falsification test in Table A.5. Columns (1), (3), and (5) present the reduced-form relationship between cotton suitability and the three outcome measures in the South, showing that the estimated effects are significant. On the other hand, columns (2), (4), and (6) show that there is no consistent relationship between cotton suitability and political attitudes outside the South. The relationship is only significant for affirmative action, but in this case the result in the opposite direction: higher cotton suitability leads to higher levels of support. If anything, such a positive relationship would bias our results in the conservative direction. As an additional test, we applied the same falsification test to a more historically complete source of data: presidential election returns (which we discuss in additional detail in Section 6.2). Drawing on county-level returns (Clubb, Flanigan and Zingale, 2006), we estimated the reduced-form relationships between cotton suitability and the percentage voting for the Democratic presidential candidate for both the South and the non-South in each presidential election from 1872 until 1972, separately. Figure A.2 plots the coefficient and

Table A.3: Effect of slavery on individual white partisanship, views on affirmative action, and racial resentment.

	Prop. Democrat			Affirm. Action			Racial Resentment		
	Logit			Logit			OLS		
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Prop. Slave, 1860	-0.582*	-0.443 [†]	-0.499*	-0.863**	-0.780**	-0.916**	0.402*	0.338*	0.401*
	(0.247)	(0.238)	(0.243)	(0.225)	(0.227)	(0.232)	(0.165)	(0.150)	(0.152)
Log(County Area)	-0.082	-0.067	-0.050	-0.198**	-0.187**	-0.161*	0.181**	0.154**	0.145*
	(0.072)	(0.070)	(0.071)	(0.071)	(0.070)	(0.072)	(0.067)	(0.057)	(0.055)
Ruggedness	0.001 [†]	0.001 [†]	0.001	-0.00001	-0.0003	-0.0002	-0.0001	0.0001	-0.000
	(0.001)	(0.001)	(0.001)	(0.001)	(0.001)	(0.001)	(0.001)	(0.0005)	(0.001)
Land Inequality, 1860	0.081	0.006	-0.093	-0.100	-0.117	-0.042	0.414	0.343	0.326*
	(0.407)	(0.425)	(0.435)	(0.359)	(0.379)	(0.430)	(0.270)	(0.274)	(0.293)
Prop. Small Farm, 1860	0.158	0.183	0.298	0.222	0.244	0.198	-0.220	-0.199	-0.14
	(0.252)	(0.256)	(0.278)	(0.250)	(0.253)	(0.274)	(0.209)	(0.191)	(0.201)
Log(Total Pop., 1860)	-0.028	-0.021	-0.036	0.007	0.025	-0.002	-0.086*	-0.063 [†]	-0.05
	(0.052)	(0.053)	(0.052)	(0.048)	(0.050)	(0.053)	(0.040)	(0.038)	(0.039)
Log(Farm Value per acre, 1860)	0.169**	0.126*	0.084	0.085	0.054	0.069	-0.037	-0.001	-0.00
	(0.061)	(0.063)	(0.065)	(0.059)	(0.059)	(0.064)	(0.043)	(0.040)	(0.042)
Prop. Free Black, 1860	0.119 [†]	0.102 [†]	0.119 [†]	0.129*	0.117*	0.133*	-0.062	-0.055	-0.06
	(0.061)	(0.060)	(0.063)	(0.056)	(0.056)	(0.060)	(0.050)	(0.045)	(0.048)
Log(Improved Acres, 1860)	3.401*	2.874 [†]	2.961 [†]	1.713	0.870	0.556	-2.548 [†]	-1.702	-1.71
	(1.693)	(1.724)	(1.730)	(1.780)	(2.009)	(2.098)	(1.323)	(1.230)	(1.268)
Rail Access, 1860	0.175**	0.135*	0.149*	0.039	0.009	0.010	-0.048	-0.008	0.000
	(0.064)	(0.062)	(0.062)	(0.063)	(0.065)	(0.066)	(0.042)	(0.039)	(0.039)
Water Access, 1860	-0.030	-0.025	-0.028	0.043	0.065	0.054	0.026	0.006	-0.00
	(0.056)	(0.059)	(0.061)	(0.060)	(0.060)	(0.062)	(0.046)	(0.043)	(0.045)
Educ: HS Grad		-0.078	-0.088		-0.366**	-0.349**		0.250**	0.296*
		(0.101)	(0.108)		(0.094)	(0.101)		(0.089)	(0.094)
Educ: Some College		0.039	0.041		-0.308**	-0.284**		0.155 [†]	0.183
		(0.105)	(0.112)		(0.100)	(0.106)		(0.088)	(0.093)
Educ: 2-year degree		-0.063	-0.075		-0.419**	-0.382**		0.324**	0.381*
		(0.122)	(0.132)		(0.116)	(0.122)		(0.098)	(0.104)
Educ: 4-year degree		0.141	0.121		-0.269**	-0.252*		-0.127	-0.07
		(0.103)	(0.109)		(0.103)	(0.109)		(0.094)	(0.099)
Educ: Postgrad		0.665**	0.650**		0.324**	0.350**		-0.460**	-0.423
		(0.121)	(0.130)		(0.119)	(0.125)		(0.093)	(0.097)
Family Inc: \$25,000-50,000		-0.008	-0.005		-0.263**	-0.270**		0.072 [†]	0.047
		(0.051)	(0.054)		(0.053)	(0.056)		(0.037)	(0.039)
Family Inc: \$50,000-100,000		-0.241**	-0.234**		-0.470**	-0.471**		0.081*	0.058*
		(0.050)	(0.052)		(0.055)	(0.058)		(0.036)	(0.038)
Family Inc: \$100,000-150,000		-0.455**	-0.414**		-0.546**	-0.531**		0.167**	0.126
		(0.069)	(0.069)		(0.074)	(0.078)		(0.051)	(0.054)
Family Inc: \$150,000+		-0.253**	-0.232*		-0.403**	-0.380**		0.114	0.086
		(0.091)	(0.094)		(0.095)	(0.102)		(0.085)	(0.089)
Religion Import.		-0.961**	-0.981**		-0.492**	-0.499**		0.395**	0.395*
		(0.047)	(0.048)		(0.048)	(0.050)		(0.039)	(0.040)
Female		0.428**	0.419**		0.382**	0.364**		-0.076**	-0.074
		(0.037)	(0.038)		(0.043)	(0.044)		(0.027)	(0.028)
Age		0.0005	0.001		-0.012**	-0.012**		0.005*	0.005*
		(0.001)	(0.001)		(0.001)	(0.001)		(0.001)	(0.001)
Prop. Black in Zip, 2000			0.109			0.351*			-0.10
			(0.186)			(0.171)			(0.126)
Log(Median Zip Code Inc, 2010)			-0.286**			-0.153 [†]			0.073
			(0.075)			(0.082)			(0.054)
White Unemp. Rate, 2010-2014			-5.522**			-1.640			1.606
			(1.451)			(1.547)			(0.865)
Log(White-Black Median Inc Ratio)		44	0.112			0.123			0.001
			(0.120)			(0.108)			(0.075)
Constant	10.224	2.644	12.457	21.158	21.789	30.542*	-17.175	-11.585	-11.4
	(14.070)	(13.952)	(14.479)	(13.425)	(13.399)	(14.274)	(11.845)	(10.890)	(11.58)
State-Year Fixed Effects	✓	✓	✓	✓	✓	✓	✓	✓	✓
Latitude/Longitude	✓	✓	✓	✓	✓	✓	✓	✓	✓
N	33,670	32,226	29,917	33,611	32,167	29,863	13,876	13,860	12,86

Note: [†]p < .1; *p < .05; **p < .01. Cluster-robust standard errors in parentheses, clustered at the county level. All models include state-year fixed effects.

Table A.4: Effect heterogeneity of slavery by individual and contextual variables.

	Prop. Democrat		Affirm. Action		Racial Resentment	
	Logit		Logit		OLS	
	(1)	(2)	(3)	(4)	(5)	(6)
Prop. Slave × Educ: HS Grad	-0.992 [†]		-0.424		-0.426	
	(0.522)		(0.511)		(0.387)	
Prop. Slave × Educ: Some College	-0.635		-0.366		-0.455	
	(0.549)		(0.550)		(0.391)	
Prop. Slave × Educ: 2-year degree	-0.428		-0.125		-0.412	
	(0.632)		(0.681)		(0.415)	
Prop. Slave × Educ: 4-year degree	-0.394		-0.565		-0.474	
	(0.561)		(0.600)		(0.399)	
Prop. Slave × Educ: Postgrad	-0.604		-0.165		-0.188	
	(0.629)		(0.635)		(0.426)	
Prop. Slave × Family Inc: \$25,000-50,000	-0.192		0.045		-0.043	
	(0.271)		(0.286)		(0.178)	
Prop. Slave × Family Inc: \$50,000-100,000	0.301		0.436		-0.132	
	(0.272)		(0.289)		(0.187)	
Prop. Slave × Family Inc: \$100,000-150,000	-0.364		-0.059		0.219	
	(0.362)		(0.389)		(0.272)	
Prop. Slave × Family Inc: \$150,000+	-0.262		-0.389		-0.170	
	(0.523)		(0.578)		(0.358)	
Prop. Slave × Religion Import.	-0.117		-0.600*		0.262	
	(0.225)		(0.254)		(0.191)	
Prop. Slave × Female	0.110		0.016		0.170	
	(0.212)		(0.237)		(0.130)	
Prop. Slave × Age	0.006		-0.0003		0.003	
	(0.006)		(0.008)		(0.005)	
Prop. Slave × Prop. Black in Zip, 2000		-3.390**		-1.304		1.370*
		(0.940)		(0.859)		(0.586)
Prop. Slave × Log(Median Zip Code Inc, 2010)		-0.545		0.211		0.289
		(0.408)		(0.419)		(0.274)
Prop. Slave × White Unemp. Rate, 2010-2014		-8.666		-9.507		-2.151
		(6.607)		(6.772)		(3.975)
Prop. Slave × Log(White-Black Median Inc Ratio)		-0.083		0.657		0.056
		(0.543)		(0.588)		(0.309)
State-Year Fixed Effects	✓	✓	✓	✓	✓	✓
Lower-Order Terms	✓	✓	✓	✓	✓	✓
Individual-Level Covariates	✓	✓	✓	✓	✓	✓
Contextual Covariates		✓		✓		✓
N	32,226	29,917	32,167	29,863	13,860	12,866

Note: [†]p < .1; *p < .05; **p < .01. Cluster-robust standard errors in parentheses, clustered at the county level. All models include state-year fixed effects, latitude/longitude and their squared terms. "Religion Import" is 1 if a respondent says that religion is important to their lives. Education omitted category is less than a high school degree. Omitted category for family income is less than \$25,000. Zip-code level data comes from the American Community Survey.

Table A.5: Reduced form relationships between cotton suitability and white attitudes in South and Non-South.

	Prop. Democrat		Affirm. Action		Racial Resentment	
	(1)	(2)	(3)	(4)	(5)	(6)
	South	Non-South	South	Non-South	South	Non-South
FAO Cotton Suitability	-0.096* (0.040)	0.010 (0.033)	-0.096** (0.034)	0.068* (0.031)	0.303* (0.135)	-0.142 (0.105)
State Fixed Effects	✓	✓	✓	✓	✓	✓
Geographic Controls	✓	✓	✓	✓	✓	✓
N	1,180	476	1,180	476	1,038	416
R ²	0.167	0.297	0.083	0.119	0.096	0.252

Note: † p < .1; * p < .05; ** p < .01. Included in the Non-South are the following states with some positive cotton suitability: AZ, CA, IA, IL, IN, KS, NE, NJ, NM, NY, OH, PA, UT.

95% confidence intervals for each of these reduced-form models and shows that there is a strong reduced-form relationship over time in the South (explored further below), but a fairly precisely estimated non-effect in the non-South for the time period between the end of the Civil War and the Voting Rights Act of 1965. Thus, both historically and today, there is little evidence that cotton suitability has any effect on attitudes in the absence of the institution of slavery, making the exclusion restriction appear reasonable in this case and lending credibility to our causal estimates.

B.3 Matching Adjacent Counties

Although our results in Table 1 control for a number of historical and geographic covariates, it remains possible that our results are driven by differences between slaveholding and non-slaveholding areas not fully captured by these covariates. For instance, it could be that the “upland” regions of northern Alabama and Georgia differed systematically from the Black Belt (as suggested by [Kousser, 2010](#)). To test the robustness of our results to such confounders, we restrict our sample to the set of neighboring counties that border a county in which proportion slaves differs by more than 20 percentage points. This enables us to compare the effects of slavery across counties that are geographically and perhaps also politically, economically, and culturally similar (as [Banerjee and Iyer, 2005](#), do with Indian districts). It also drops certain former high slave counties that are in regions where all of the neighbors are also high slave areas—for example, the Mississippi Delta. Columns (1), (3), and (5) of Table A.6 show that the results for all three of our CCES outcomes are robust to restricting our analysis to only these neighboring counties, even though this removes more than half of the counties in our original sample. Thus, even within fairly geographically concentrated areas, there

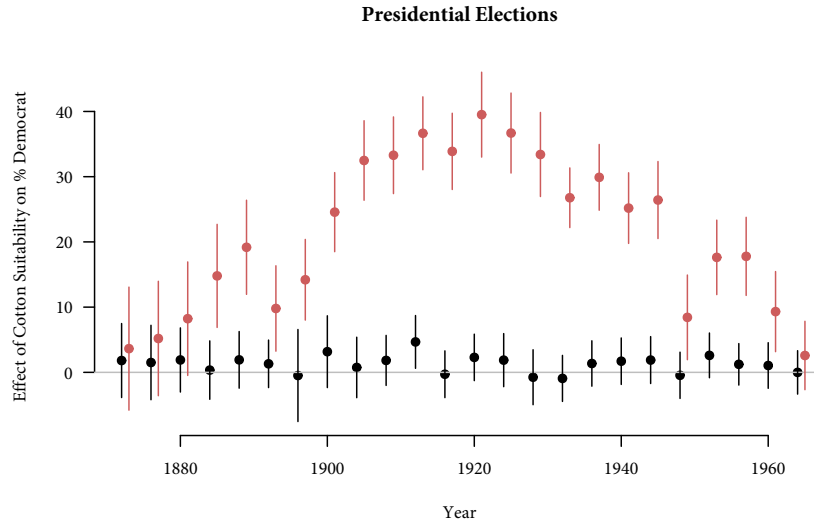


Figure A.2: Reduced-form coefficients and 95% confidence intervals for the effect of cotton suitability on the county Democratic vote-share in presidential elections in the South (red) and the non-South (black).

are strong, statistically significant differences between counties with higher and lower past concentrations of slaves.

B.4 Counterfactual Comparisons to the North

If the effects that we estimate in Table 1 are genuinely attributable to the local prevalence of slavery, then we should see no difference in our outcomes between areas of the South that were largely non-slaveholding and areas in other parts of the country that also did not have slaves, such as counties in the North. In addition, if no such differences exist, then that would provide evidence against the alternative theory that it is the institutional *legality* of slaveholding, rather than the *local prevalence* of slavery, that is driving our results. Making these comparisons with the North also enables us to address what we consider to be the appropriate counterfactual, which is what contemporary political attitudes in the South would have been had slavery been as non-prevalent in the South as it was in the North.

We therefore examine differences between Southern counties with very few slaves in 1860 and non-Southern counties with no slaves in 1860. To do this, we restrict the data to counties in slave states where fewer than 5% of the county population was enslaved, and then match these counties to similar counties in non-slaves states on geography (latitude/longitude), county size, farm value per capita, mixed-race population,

Table A.6: Neighbor matching within South and between South and Non-South.

	Prop. Democrat		Affirm. Action		Racial Resentment	
	(1)	(2)	(3)	(4)	(5)	(6)
Prop. Slave, 1860	-0.202** (0.068)		-0.149** (0.056)		0.450* (0.228)	
Slave State		0.015 (0.030)		0.017 (0.029)		-0.068 (0.103)
State Fixed Effects	✓		✓		✓	
1860 Covariates	✓	✓	✓	✓	✓	✓
50% Threshold Match	✓		✓		✓	
North-South Match		✓		✓		✓
N	444	384	444	384	393	354
R ²	0.248	0.118	0.156	0.077	0.195	0.106

Note: †p < .1; *p < .05; **p < .01. Columns (1), (3) and (5) show results of regressions with state fixed effects and 1860 covariates for those counties that border a county in which proportion slave is more than 20 percentage points different. Columns (2), (4) and (6) show difference between slave-state counties with few slaves (< 5% of 1860 population) and non-Southern counties, matched on geography, farm value per capita, and total population. Coefficients are from a regressions on the matched data, that include a dummy variable for “slave state” as well as the 1860 covariates. All models are WLS with county sample sizes as weights.

and total county population.³¹ We regress each of our three CCES outcome variables on the 1860 covariates as well as on a dummy variable for the county being in a slave state. Columns (2), (4) and (6) of Table A.6 show these results and confirm that there exists no difference between the Southern counties and the non-Southern counties beyond the effect of local slave prevalence. This provides evidence that the local prevalence of slavery, rather than state laws permitting the ownership of slaves, drives our results.

B.5 Rural versus Urban Counties

One possible explanation for our findings not covered in 5 is that large slaveholding counties tend to be more rural today than counties that have smaller slave proportions, maybe because they had plantations and other large farms. Our results might therefore reflect the simple fact that rural counties tend to be more conservative than urban counties.

To examine this possibility, we remove from our dataset the top ten percentile of all Southern counties in terms of 1860 population density. Thus, we remove ten percent

³¹We use coarsened exact matching (CEM) on these variables, employing the default cut-points (Iacus, King and Porro, 2012; Stefano Iacus and Gary King and Giuseppe Porro, 2009). To avoid biasing our results, we drop Maryland and Missouri from the Northern sample since both had non-trivial slavery in 1860. Replicating this analysis with propensity score matching or genetic matching does not substantively change the results (available upon request).

Table A.7: The effects of slavery after eliminating large urban centers and controlling for Civil War destruction.

	Prop. Democrat		Affirm. Action		Racial Resentment	
	(1)	(2)	(3)	(4)	(5)	(6)
Prop. Slave, 1860	-0.097*	-0.124**	-0.131**	-0.173**	0.346*	0.487**
	(0.047)	(0.044)	(0.041)	(0.038)	(0.154)	(0.144)
Civil War Destruction		0.001		-0.006		0.025
		(0.008)		(0.006)		(0.027)
State Fixed Effects	✓	✓	✓	✓	✓	✓
1860 Covariates	✓	✓	✓	✓	✓	✓
Dense 1860 Counties Dropped	✓		✓		✓	
N	1,030	1,011	1,030	1,011	909	909
R ²	0.181	0.198	0.082	0.098	0.089	0.137

Note: †p < .1; *p < .05; **p < .01. All models are WLS with within-county sample size as weights. Dense counties are those defined to be in the top ten percentile of population density in 1860.

of counties that have been historically the most urban. Removing these counties hardly changes the estimated effects of slavery, as indicated in columns (1), (3) and (5) of Table A.7. Our results are therefore unlikely to be attributable to the sparse populations of former slaveholding counties.³²

B.6 Civil War Destruction

Another possible explanation for our findings not covered in Section 5 is that slaveholding counties were more adversely affected by the Civil War (1861–1865). The damage to infrastructure and the loss of life resulting from the War was extensive and affected the South’s agricultural areas disproportionately (Goldin and Lewis, 1975).³³ This could affect our analysis in two ways. First, in light of the federal government’s role in the war, whites in war-torn slave counties perhaps became more resentful of the federal government, which in turn they express through resentment toward blacks. Second, it may be that the Civil War disrupted the social fabric of these communities, aggravating racial strife in the process (Collier et al., 2013).

In either case, we would expect the effects of slavery to diminish once we control for Civil War destruction. Columns (2), (4) and (6) of Table A.7 therefore control for

³²An alternative strategy is to restrict our attention to counties that are rural today as opposed to rural counties in 1860. This approach, however, potentially suffers from post-treatment bias since the prevalence of slavery in 1860 could affect population density today. Mindful of this possibility, we include such an analysis in the Appendix, noting that the results are consistent with those of Table A.7. Similar results are obtained when controlling for modern-day county population as opposed to 1860 county population.

³³The correlation between proportion slave in 1860 and our measure of the Civil War’s impact (described below) is positive at 0.23.

Table A.8: The effect of slavery and antebellum attitudes.

	Prop. Democrat		Affirm. Action		Racial Resentment	
	(1)	(2)	(3)	(4)	(5)	(6)
Prop. Slave, 1860	-0.226** (0.063)	-0.118** (0.046)	-0.260** (0.058)	-0.179** (0.040)	0.828** (0.213)	0.499** (0.152)
Log Relative Slave Mortality, 1860	0.050** (0.015)		0.009 (0.014)		-0.132* (0.051)	
Avg. Residents per Slave Dwelling		-0.012** (0.003)		-0.006* (0.003)		0.036** (0.010)
State Fixed Effects	✓	✓	✓	✓	✓	✓
1860 Covariates	✓	✓	✓	✓	✓	✓
N	495	885	495	885	447	795
R ²	0.251	0.232	0.146	0.106	0.258	0.145

Note: † p < .1; * p < .05; ** p < .01. All models are WLS with within-county sample size as weights. Log Relative Slave Mortality is the natural log of the ratio of the slave mortality rate and the white mortality rate, both measured in 1859. Average Residents per Slave Dwelling were calculated by dividing the number of slaves on a farm, divided by the number of slave dwellings, and then average across farms in the same county.

U.S. Census measurements of the percentage drop in the average value of farms in the county between 1860 and 1870, which is a proxy for Civil War destruction.³⁴ As Table A.7 indicates, slavery's effects on our three outcome measures are hardly affected by the inclusion of this variable.³⁵ Furthermore, in results not presented here, we find that even when we include an interaction term between proportion slave and Civil War destruction, the interaction is not significant.

C Supplemental Information: Antebellum Measures

Can antebellum attitudes of whites completely explain our results? We addressed this question in the main text by examining presidential vote shares. Here, we provide some additional evidence to shed light on this question.

First, we attempt to proxy the antebellum attitudes of whites toward slaves by mea-

³⁴We assume that ignorability is satisfied here for both slavery and Civil War destruction with the same set of covariates, which would make the effect on the slave variable the controlled direct effect. We believe that this is a more plausible assumption than that made with respect to contemporary black population. The reason is because conditioning on 1860s covariates is likely to result in an accurate estimate of which counties suffered more destruction during the Civil War, but would probably not be sufficient to separately identify the effect of black concentrations in the 20th century.

³⁵Since the nature of land value changed dramatically before and after the Civil War due to the emancipation of slaves (Wright, 1986), we use an alternative measure of destruction based on the loss in livestock value in Table A.13. Results using this measure are almost identical to those presented here.

asuring the relative mortality of slaves to whites, as measured by the 1860 Census.³⁶ To account for between county variation in mortality, we measure the relative mortality of slaves to free whites in a county using the natural log of the ratio of the slave mortality rate to the white mortality rate. The mortality rate here refers to the number of deaths divided by the size of the subpopulation in 1860. Negative racial attitudes could have led white planters, farmers, or overseers to increase the mortality of slaves, either directly through violence or indirectly through overwork, undernourishment, and poor medical care. Thus, relative slave mortality may, in part, reflect how whites viewed and valued black lives in the antebellum period. If the effect of slavery disappears when controlling for relative slave mortality in our baseline model, we might conclude that antebellum attitudes are driving our findings. These results, reported in columns (1), (3), and (5) of Table A.8, are largely consistent with our baseline models, albeit with slight differences due to the availability of the mortality data. While these models represent an imperfect test, they do suggest that antebellum attitudes cannot explain the entirety of our baseline results.

The second is to explore the local treatment of slaves as a proxy for attitudes about race. Comprehensive data on racial views are not available in the antebellum period, so we instead look for measures that might be consequences of such attitudes. In particular we use samples from the slave schedules of the 1860 U.S. Census to calculate the average occupant size of slave quarters on farms in a county (Menard et al., 2004). Across the South, the average slave quarters housed around five individuals, though this number varied considerably across counties. Variation in the occupancy of such quarters may come from both variation in the size of slave families and also the propensity of farm owners to place multiple families in the same dwelling. Attitudes about race might affect both of these sources. First, there is some evidence that planters engaged in so-called “slave breeding,” which entailed various ways of promoting and forcing high fertility rates among enslaved women (Sutch, 1975), though the extent of this practice is contested (Fogel and Engerman, 1995). Second, those planters with more extreme negative racial attitudes might provide less housing for their slaves, which would be measured as a higher occupancy in the average slave dwelling. Obviously there are other factors that affect this measure, but it may pick up some degree of planter cruelty or racial animus that is not captured by the density of slavery alone.

We add the the average number of occupants of slave quarters to our baseline specification in columns (2), (4), and (6) in Table A.8. Here we see that both the economic institution of slavery, as measured by proportion slave, and the relative treatment of slaves, as measured by the dwelling size, have independent and significant effects on

³⁶We coded the mortality information from the original manuscripts of the Mortality Schedules of the 1860 U.S. Census, which asked census takers to record all deaths that occurred in the year 1859. Data was also collected on the enslaved status, occupation, and age of the deceased. Manuscripts were only available for Alabama, Georgia, Kentucky, Louisiana, South Carolina, Tennessee, Texas, and Virginia.

the attitudes of whites today. Despite the potential significant effect of slave dwelling size on the outcome variables, however, we still see a strong effect of proportion slave on attitudes as well, indicating that the localized prevalence of slavery continues to matter once we account for aspects of how slaves were treated.

D Supplemental Information: Additional Evidence on Geographic Sorting

We also present some suggestive evidence against geographic sorting using data on between-county migration from 1995 to 2000 from the 2000 U.S. Census (Bureau, 2001). These data help us investigate the extent to which *contemporary*, as opposed to historical migration, explains our findings (see Dell, 2010, for a similar analysis). In order for geographic sorting to explain our results, two conditions must hold. The first is that there must be migration from low-slave areas to high-slave areas (or vice-versa); otherwise, there is no meaningful sorting of any kind. To test this condition, we use county-to-county migration data to calculate dyads of where people move to and from, measuring the absolute difference in the proportion 1860 slave between the departing and receiving county; this enables us to assess how much migration exists between low-slave and high-slave areas. Appendix Figure A.3 shows the relationship between these flows and the difference in proportion slave, and it demonstrates that, as the slavery differential grows, the migrations between counties drops significantly. Thus, the vast majority of contemporary migration is within low-slave areas or within high-slave areas, not between.

The second condition for sorting to explain our findings is that racially conservative whites are moving into high-slave areas, racially liberal whites are moving out of high-slave areas, or some combination. Even if there is very little migration between low- and high-slave counties (as shown in Appendix Figure A.3), the distribution of political beliefs among these migrants could be so highly skewed so as to produce our results. (For example, perhaps all of the out-migrants from high-slave counties are racially liberal and all of the in-migrants to high-slave counties are racially conservative.) With regard to the first possibility, this seems unlikely to be the primary mechanism as it relies on racially hostile whites moving to areas with extremely large proportions of African Americans. For example, Farley et al. (1994), show that anti-black attitudes are correlated with stronger preferences for geographic segregation. More plausible is the exodus of racially liberal whites from former slave counties. To check this, again using contemporary data, we examine the relationship between the proportion slave in 1860 and out-migration in 1995-2000 census records. We find that proportion slave actually has a *negative* effect on contemporary out-migration (Appendix Table A.9). Thus, we have no evidence drawn from contemporary data for any of the necessary conditions

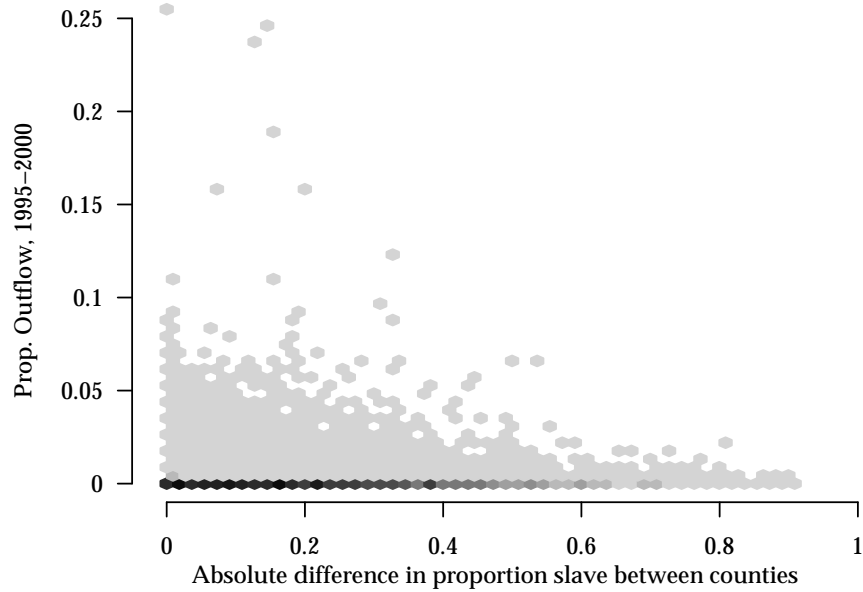


Figure A.3: Relationship between county-to-county migration and the similarity of those counties on proportion slave. The darker the hexagon, the more county-dyads in that bin. The x-axis represents the absolute difference between the counties in term of proportion slave in 1860 and the y-axis represents the migration (1995–2000) from the sending county to the receiving county as a proportion of the sending county’s 2000 population.

behind a geographic sorting explanation.

E Supplemental Information: Additional Tables and Figures

Table A.9: Effect of slavery present-day migration.

	In-migration (1)	Out-migration (2)
Prop. Slave, 1860	-0.055** (0.015)	-0.055** (0.015)
State Fixed Effects	✓	✓
1860 Covariates	✓	✓
N	1,207	1,207
R ²	0.183	0.184

Note: †p < .1; *p < .05; **p < .01. Out-migration is defined as the total number of out-migrants from a county from 1995 to 2000 divided by the total population in 2000. In-migration is similarly defined.

Table A.10: Feeling thermometer score effect disaggregated by race.

	Black Therm. Scores		White Therm. Scores	
	OLS	IV	OLS	IV
	(1)	(2)	(3)	(4)
Prop. Slave, 1860	-18.678* (7.710)	-19.261** (5.043)	10.452 (6.847)	17.226† (9.494)
Clustered SEs	✓	✓	✓	✓
State-Year Fixed Effects	✓	✓	✓	✓
Geographic Controls	✓	✓	✓	✓
1860 Covariates	✓		✓	
N	2,407	2,006	1,517	1,277
R ²	0.112	0.069	0.085	0.069

†p < .1; *p < .05; **p < .01. All analyses are at the individual level with standard errors clustered at the county level. Data from the ANES 1984-1998.

Table A.11: County-level regression models for the effect of slavery on difference between ANES feeling thermometer scores for whites and blacks.

	White Thermometer - Black Thermometer (-100 to 100)			
	OLS			IV
	(1)	(2)	(3)	(4)
Prop. Slave, 1860	27.215** (7.379)	19.850** (6.735)	28.856 [†] (14.358)	51.748* (19.305)
Clustered SEs	✓	✓	✓	✓
State-Year Fixed Effects	✓	✓	✓	✓
Geographic Controls	✓	✓	✓	✓
1860 Covariates	✓		✓	
N	56	56	54	51
R ²	0.201	0.715	0.793	0.675

[†]p < .1; *p < .05; **p < .01. All analyses are at the county level, weighted by within-county sample sizes adjusted for ANES survey weights. Data from the ANES 1984-1998.

Table A.12: First stage estimates from the sequential g-estimation model of Table 3.

	Prop Democrat (1)	Affirm. Action (2)	Racial Resentment (3)
Prop. Slave, 1860	-0.092 [†] (0.048)	-0.080* (0.041)	0.289 [†] (0.156)
Prop. Black 2000	0.028 (0.056)	-0.112* (0.048)	-0.072 (0.179)
Log Population, 2000	0.017** (0.006)	0.023** (0.005)	-0.071** (0.021)
Percent High School Graduates, 1990	0.003** (0.001)	0.002* (0.001)	-0.006 [†] (0.003)
Unemployment, 1999	0.004 (0.003)	-0.003 (0.003)	0.023* (0.011)
Median Income, 2000	-0.139** (0.042)	-0.208** (0.036)	0.460** (0.136)
White-Black Income Ratio, 2000	0.016 (0.012)	0.013 (0.010)	-0.019 (0.039)
State Fixed Effects	✓	✓	✓
1860 Covariates	✓	✓	✓
N	1,093	1,093	976
R ²	0.241	0.148	0.186

Note: [†]p < .1; *p < .05; **p < .01. Note that conditional on the past, proportion black today lacks explanatory power. While the estimates of proportion slave are insignificant in these models, their estimates possess large amounts of post-treatment bias due to the contemporary variables. Each model includes weights for the within-county sample size.

Table A.13: Alternative measure of civil war destruction based on livestock value.

	Prop. Democrat (1)	Affirm. Action (2)	Racial Resentment (3)
Prop. Slave, 1860	-0.113* (0.045)	-0.171** (0.039)	0.427** (0.146)
Livestock Value Loss, 1860-1870	0.0003 (0.001)	0.0005 (0.001)	-0.002 (0.004)
State Fixed Effects	✓	✓	✓
1860 Covariates	✓	✓	✓
N	1,010	1,010	908
R ²	0.198	0.097	0.137

Note: † p < .1; * p < .05; ** p < .01. Livestock value loss is the decrease in the value of livestock in a county between 1860 and 1870 as a proportion of 1860 livestock value.

Table A.14: Effect of slavery on measures of inequality and income.

	log White-Black Wage Gap, 1940		log White-Black Med. Income Gap, 2014		log White Median Income, 2014	
	OLS	IV	OLS	IV	OLS	IV
	(1)	(2)	(3)	(4)	(5)	(6)
Prop. Slave, 1860	0.578** (0.185)	1.355** (0.478)	0.414** (0.095)	1.065** (0.215)	0.134* (0.053)	0.852** (0.126)
State Fixed Effects	✓	✓	✓	✓	✓	✓
Geographic Controls	✓	✓	✓	✓	✓	✓
1860 Covariates	✓		✓		✓	
N	877	866	994	1,020	1,152	1,120
R ²	0.188	0.150	0.147	0.048	0.444	0.239

Note: † p < .1; ** p < .05; *** p < .01. Outcomes from the 1940 U.S. Census and the 2009-2014 American Community Survey.

Table A.15: Interaction of slavery with income on individual white attitudes.

	Prop Democrat	Affirm. Action	Racial Resentment
	logistic	logistic	OLS
	(1)	(2)	(3)
Prop. Slave, 1860	-0.025 (1.205)	-2.239 [†] (1.196)	0.446 (0.831)
Income	-0.048 (0.059)	-0.268** (0.037)	-0.072** (0.024)
Prop. Slave × Income	-0.034 (0.112)	0.160 (0.112)	-0.015 (0.076)
State Fixed Effects	✓	✓	✓
1860 Covariates	✓	✓	✓
Clustered SEs	✓	✓	✓
N	29,900	29,874	12,199
R ²			0.032
AIC	35,613.010	32,054.980	

Note: [†]p < .1; *p < .05; **p < .01.

Table A.16: How the effect of slavery varies by declines in the black population in the 20th century.

	Prop. Democrat		Affirm. Action		Racial Resentment	
	(1)	(2)	(3)	(4)	(5)	(6)
Prop. Slave, 1860	-0.086 [†] (0.050)	-0.060 (0.054)	-0.116** (0.043)	-0.122** (0.047)	0.284 [†] (0.164)	0.097 (0.178)
Prop Black Decline, 1940-1920	0.076 (0.268)		0.044 (0.230)		-1.093 (0.852)	
Prop Black Decline, 1970-1920		-0.115 (0.109)		-0.156 [†] (0.094)		0.381 (0.335)
Prop Slave × Black Decline, 1940-1920	-0.499 (0.598)		-0.482 (0.512)		3.152 [†] (1.898)	
Prop Slave × Black Decline, 1970-1920		-0.063 (0.224)		0.119 (0.192)		0.712 (0.692)
State Fixed Effects	✓	✓	✓	✓	✓	✓
1860 Covariates	✓	✓	✓	✓	✓	✓
N	1,152	1,151	1,152	1,151	1,027	1,026
R ²	0.204	0.206	0.103	0.104	0.146	0.152

Note: [†]p < .1; *p < .05; **p < .01.

Table A.17: Effect of slavery on intermediate outcomes.

	Black-White Tenancy Share Gap, 1925		Black-White Owner Share Gap, 1925	
	OLS	IV	OLS	IV
	(1)	(2)	(3)	(4)
Prop. Slave, 1860	0.102* (0.048)	0.894** (0.153)	-0.029 (0.048)	-0.700** (0.144)
State Fixed Effects	✓	✓	✓	✓
Geographic Controls	✓	✓	✓	✓
1860 Covariates	✓		✓	
N	1,029	1,005	1,029	1,005
R ²	0.372	0.046	0.325	0.105

Note: [†]p < .1; *p < .05; **p < .01. Black-White Tenancy Share Gap is the difference between the proportion of all black farms under tenancy agreements minus the proportion of all white farms under tenancy agreements. Black-White Owner Share Gap is the difference between the proportion of all black farms owned by the (black) operator minus the proportion of all white farms owned by the (white) operator.

Table A.18: Association between tractor growth and racial disparities.

	Lynchings per 100,000 residents	Log Black-White Wage Ratio, 1940
	(1)	(2)
Tractor Growth, 1930-1940	11.281 (17.177)	-1.021 (0.818)
Prop. Slave, 1860	14.165** (4.220)	-0.602** (0.185)
State-Year Fixed Effects	✓	✓
1860 Covariates	✓	✓
N	1,145	871
R ²	0.391	0.190

Note: †p < .1; *p < .05; **p < .01.